



Time series analysis and examples

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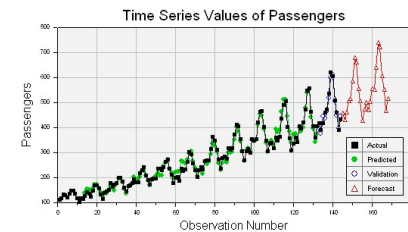
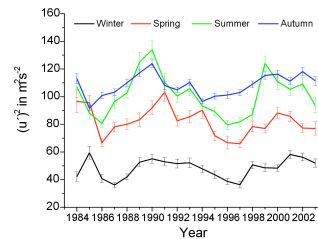
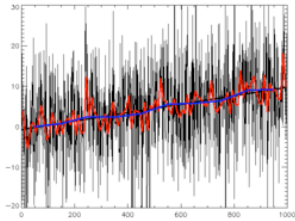


Outline

- **Introduction**
- **Preprocessing (trend, seasonality)**
- **Frequency analysis for stationary signals**
 - Model-free approach
 - Model-based approach
- **Empirical Orthogonal Functions (EOF)**
- **Non stationary signals**
 - Time-Frequency analysis
 - Time-Scale analysis
 - Wigner-Ville analysis
- **Hidden Markov models**

■ Time series

- Data obtained from consecutive measurements at equally spaced time intervals



<http://www.statsoft.com/textbook/time-series-analysis>

time series of seasonal mean gravity wave activity
<http://www.itl.nist.gov/div898/handbook/pmc/section4/pmc41.htm>

Monthly international passenger totals in 12 years
Box&Jenkins 1976

■ Extraction of information for

- Understanding of the system producing the data
- Forecasting, monitoring or control
- Identification of patterns in random signals

■ Time series analysis

- Model-free approach
- Model-based approach
- Time domain
- Frequency domain



Introduction

■ Applications

- Economic and sales Forecasting
- Stock Market Analysis
- Process and Quality Control
- Signal analysis (seismic, biological, speech, audio,..)

■ Examples in Environmental sciences

- Trends in temperature series
- Solar radiation measured by a pyranometer
- Wind speed time series
- Sea level records
- Internal gravity waves
- Seismic pattern dynamics
- Detection of cycles in atmospheric, geomagnetic and solar data

■ Time series and stochastic process

- Measured data : realizations of a more general stochastic process
- Stochastic process : random variable $x(t)$ depending on time
- Time series : discretized stochastic process

- sampling period : T

- $T_k = kT$

- Conventionally $T=1$

- Univariate time series

- Means a n -dimensional vector $\mathbf{x} = (x(1) \ x(2) \ \dots \ x(n))^T$

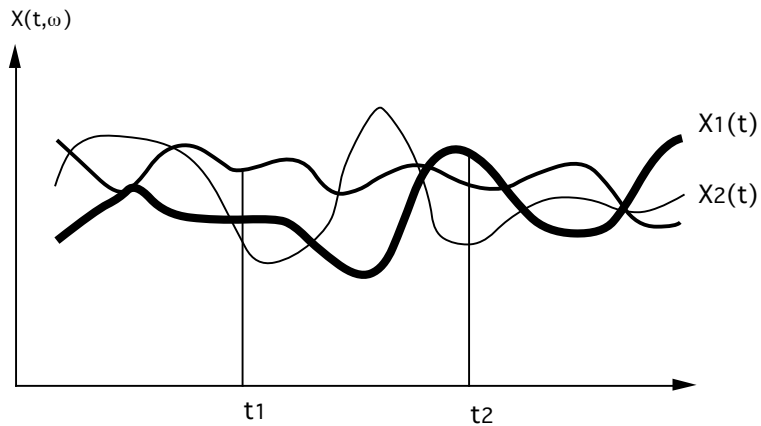
- Multivariate time series

- Each data is itself a m -dimensional vector (m simultaneous measurements)

$$\mathbf{x}(k) = (x_1(k) \ x_2(k) \ \dots \ x_m(k))^T$$

- Means a $n \times m$ dimensional matrix $\mathbf{X} = (\mathbf{x}(1) \ \mathbf{x}(2) \ \dots \ \mathbf{x}(n))^T$

■ Time series and stochastic processes

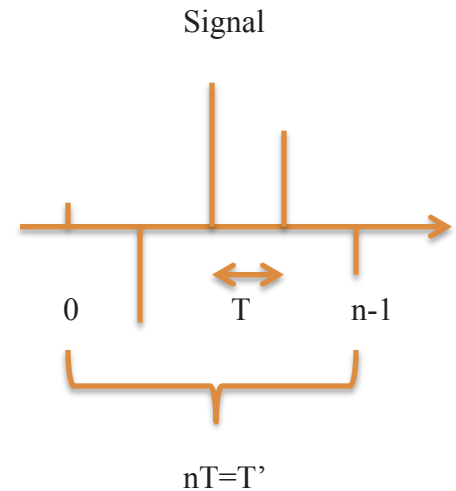


Sampling theorem



$$T \geq \frac{1}{2f_M}$$

f_M : maximal frequency



Time series

Continuous process
(many realizations)

Stochastic analysis of time series

■ Real continuous random variable

- Probability density $p(x)$

- Moments $\mu_x^{(k)} = E[(x - E(x))^k] = \int_{-\infty}^{\infty} (x - E(x))^k p(x) dx$

- Mean $m_x = E[x] = \int_{-\infty}^{\infty} xp(x) dx$

- Variance $\sigma^2 = E[(x - E[x])^2] = \int (x - E[x])^2 p(x) dx$

■ Stochastic process

- Probability density $p(\mathbf{x}) = p(x(1), x(2), \dots, x(n))$

- Moments

- Mean (order 1) $E(x(t_k)) = \int x(t_k) p(x(t_k)) dx(t_k)$

- Autocovariance (order 2)

$$\Gamma(x(t_k), x(t_l)) = \iint (x(t_k) - E(x(t_k)))(x(t_l) - E(x(t_l))) p(x(t_k), x(t_l)) dx(t_k) dx(t_l)$$

- Order k $E[(x(t_1) - E(x(t_1)))(x(t_2) - E(x(2))) \dots \dots \dots (x(t_n) - E(x(t_n)))]$

Characteristic features

- Stationarity (first and second order)

- Mean : independent of time k $m_y(k) = m_y \quad \forall k$
- Variance : independent of k $\Gamma(0) = E[(x(k) - E[x(k)])^2]$
- Covariance : depends only on the interval $l - l' = k$

$$\Gamma(k) = E[(x(l) - E(x(l))) (x(l+k) - E(x(l+k)))]$$

- Linearity

- Data are obtained by linear combination of white noise and past values (parametric model)

Characteristic features

- Trend

- Increase or decrease of the mean  non stationary (order1)

- Seasonality

- Periodically fluctuating pattern
- Relationship with cycles of the real life (hour, day, month, year,..)

- Ergodicity

- Equivalence between

- statistical mean ($E[\]$)

- Empirical mean $\lim_{n \rightarrow \infty} \frac{1}{n} \sum_{k=1}^n x(k)$ computed from a single realization on infinite observation duration $T' = nT$

Estimation of characteristic features (stationary case)

- Empirical mean

$$\hat{m} = \frac{1}{n} \sum_{k=1}^n x(k)$$

- Estimated variance

$$\hat{\sigma}^2 = \frac{1}{n} \sum_{k=1}^n (x(k) - \hat{m})^2$$

- Estimated covariance (or correlation with nul mean)

$$\hat{\Gamma}(k) = \frac{1}{n} \sum_{l=1}^n (x(l) - \hat{m})(x(l+k) - \hat{m})$$

Preprocessing



■ Trend analysis and removal

- Smoothing for noise and outliers elimination

- Fitting a function by least square

$$\min_{\theta} \sum_{k=1}^n e_{\theta}^2(k)$$

- Linear

$$e_{\theta}(k) = x(k) - (\alpha + \beta k) \quad \theta = (\alpha, \beta)$$

- Exponential

$$e_{\theta}(k) = x(k) - \exp(\alpha + \beta k) \quad \theta = (\alpha, \beta)$$

- Polynomial

$$e_{\theta}(k) = x(k) - (\alpha + \beta k + \gamma k^2) \quad \theta = (\alpha, \beta, \gamma)$$

Preprocessing



■ Seasonal analysis and removal

- Computation of Autocovariance in a specified range of lags
- Detection of the periodicity between two peaks of autocovariance
- Substraction (or division) of the seasonal component from the time series (depending of the additive or multiplicative model)

Frequency domain analysis

■ Change of Information representation system

Time: $x(t)=a\sin(2\pi ft+\varphi)$ Frequency: a, f, φ

- Well fitted for oscillating or cyclic signals, but can be applied to any signal of finite energy.

■ Tool : Fourier Transform

- Signals are represented as a weighted sum of sine and cosine waves with different frequencies (inverse transform)

$$x(t) = \int_{-\infty}^{\infty} X(f) e^{j2\pi ft} df$$

- It gives the spectrum of the signal

$$X(f) = \int_{-\infty}^{\infty} x(t) e^{-j2\pi ft} dt$$

Frequency domain analysis

■ Application to stochastic processes

- The Fourier transform of a realization of a stochastic processes may not exist, because integration can diverge.
- Power spectral density : Fourier transform of the autocovariance

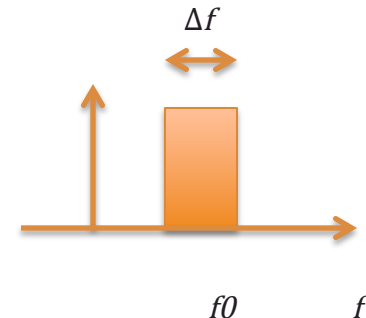
$$S(f) = \int \Gamma(\tau) e^{-j2\pi f\tau} d\tau$$

- Physical meaning



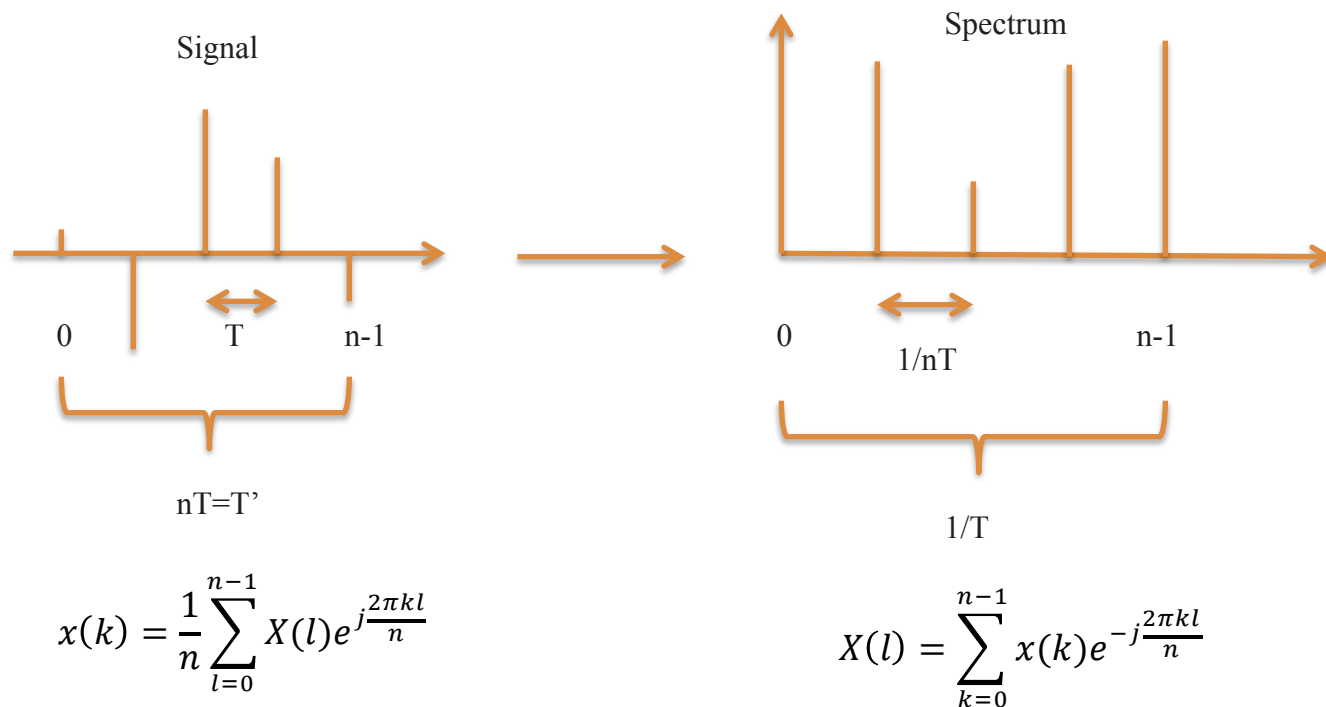
- Instantaneous power $E(y_f^2(t))$

- density $S_x(f) = \lim_{\Delta f \rightarrow 0} \left[\frac{1}{\Delta f} E(y_f^2(t)) \right]$



Frequency domain analysis

- Application to discretized signals
 - Spectrum (Discrete Fourier Transform)



Frequency domain analysis

■ Power spectrum estimation of time series

- Applied to stationary time series (TS)
- From a single realization (experiment)
- With limited time record

■ Quality of estimation

- Resolution (depends on the TS length)
- Variance of the estimator (accuracy)
- Dynamic



Tradeoff between resolution and accuracy

Frequency domain analysis

■ True power spectrum

$$S_x(k) = \sum_{l=0}^{n-1} \Gamma(l) e^{-j \frac{2\pi l k}{n}} \quad k = 0, \dots, n-1$$

■ Estimated power spectrum : Periodogram

• Two equivalent methods

- Estimation of covariance and discrete Fourier transform of covariance

$$\Gamma(k) = \frac{1}{n} \sum_{l=0}^{n-1} x(l)x(l+k)$$

$$S(k) = \sum_{l=0}^{n-1} \Gamma(l) e^{-j \frac{2\pi l k}{n}} \quad k = 0, \dots, n-1$$

- Discrete Fourier transform of TS and square modulus

$$X(f) = \sum_{k=0}^{n-1} x(k) e^{-j 2\pi k f}$$

$$S_x(f) = \frac{1}{n} |X(f)|^2$$

Frequency domain analysis

■ Periodogram : bad estimator

- Accuracy independent of the number n of data
- Increase of n improves only resolution

■ Needs for accuracy improvement

- Averaging (Welch's method)
 - Division of TS in N frames
 - Computation of periodogram on each frame
 - Mean of periodograms
- Filtering (Bartlett's method)
 - Discrete Fourier transform of windowed estimated correlation

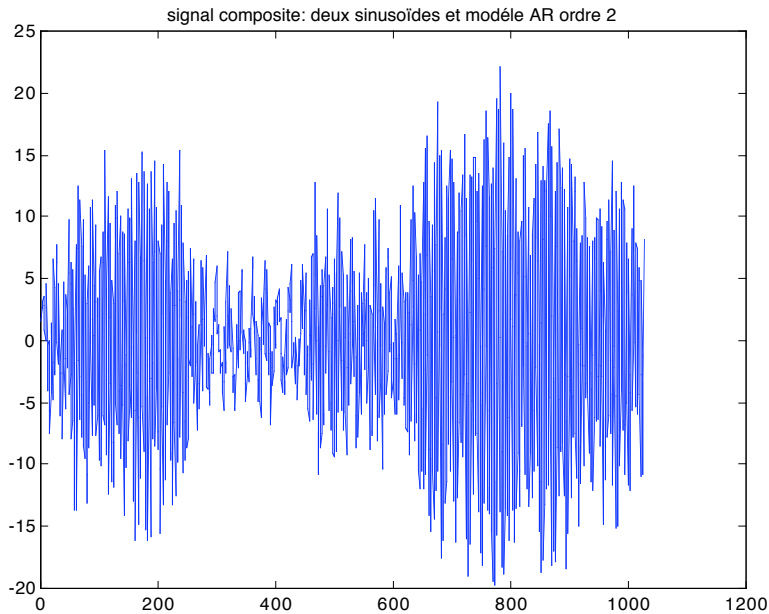


Increase of accuracy, decrease of resolution

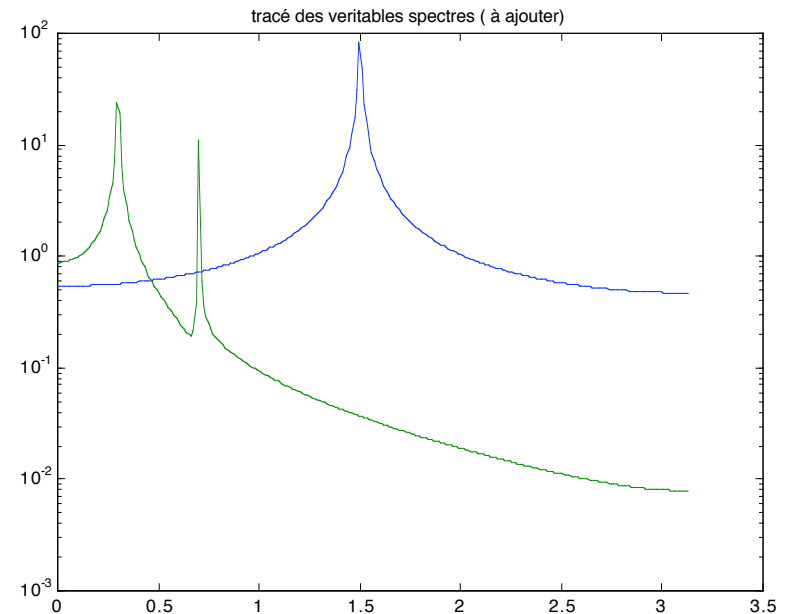
Frequency domain analysis

■ Example

- Complex signal as a sum of two sine waves and a stationary stochastic process with a damped correlation

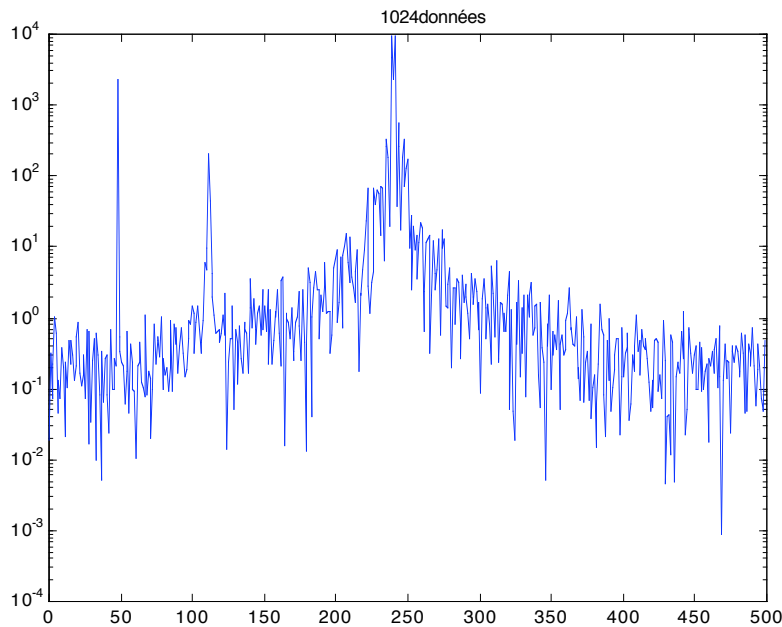


Signal

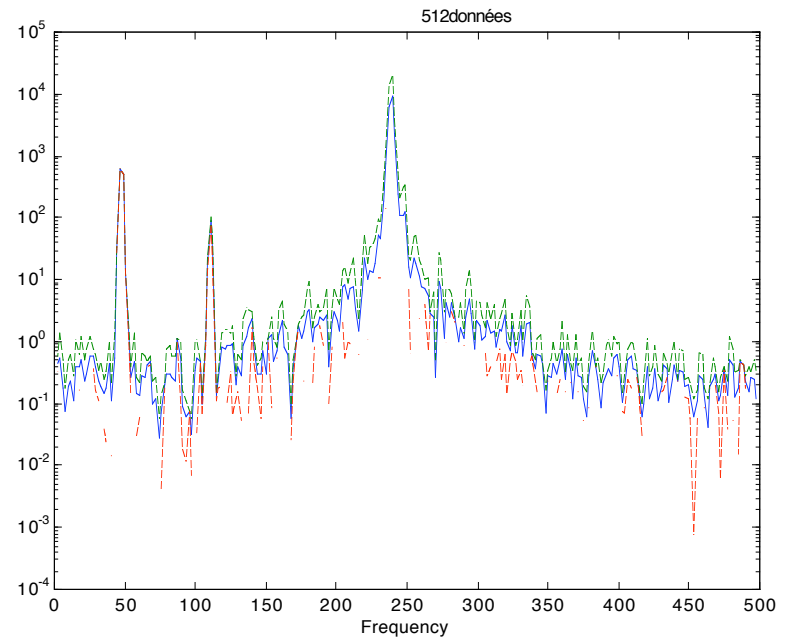


Theoretical power spectrum

Frequency domain analysis

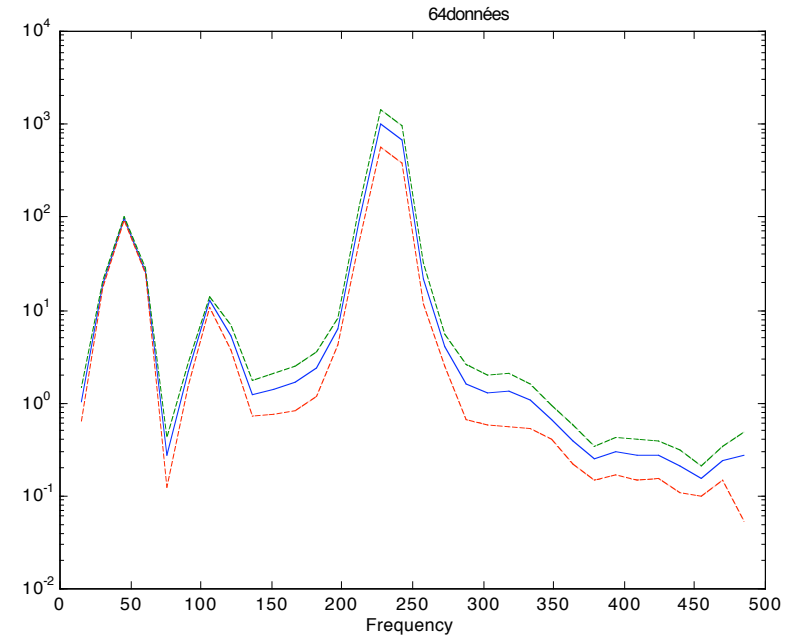
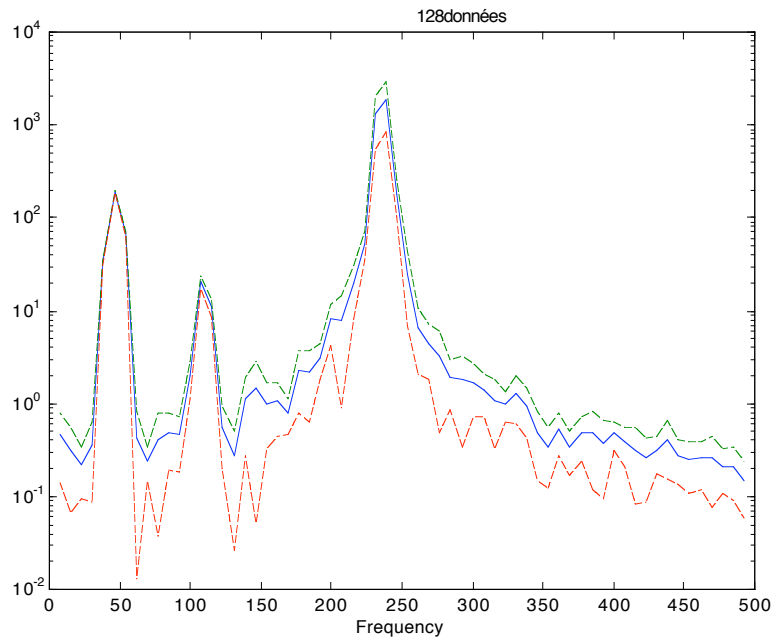


■ Periodogram 1024 data



■ Improved periodogram 512 data

Frequency domain analysis



■ Improved periodogram 128 data



■ Improved periodogram 64 data

Frequency domain analysis

■ Advantages

- Usual
- Show periodicity, cycles
- Fits with stationary signals

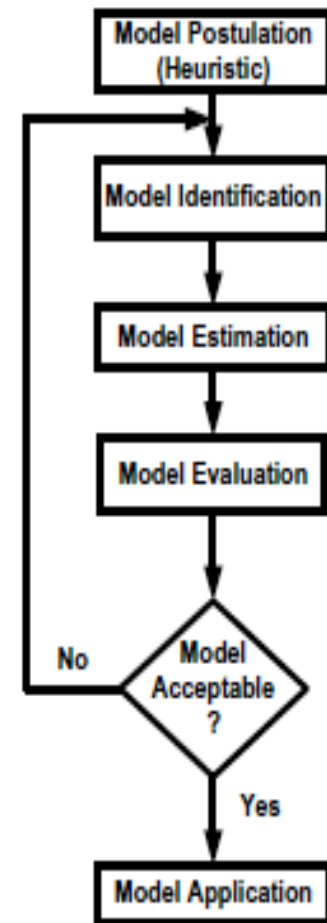
■ Drawbacks

- Does not show discontinuities, change of frequencies  non stationary methods
- Does not fit with short data length  model-based approach

Parametric model

■ Model building (*Box and Jenkins, 1976*)

- Model identification
 - number of model parameters
- Model estimation
 - Estimation of the value of the parameters
- Model validation
 - Checking of the model accuracy
- Model forecasting
 - Use of the model to establish the confidence limits of the forecast



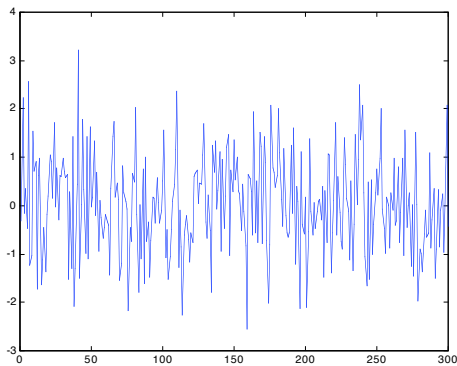
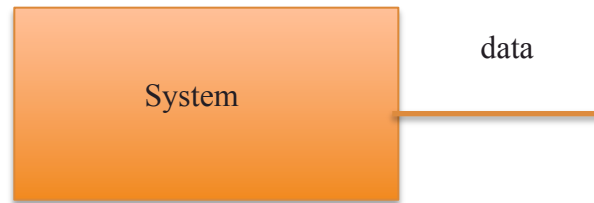
Parametric model

■ Simplest model for systems producing data

- Linear
- Stationary



Filter definition

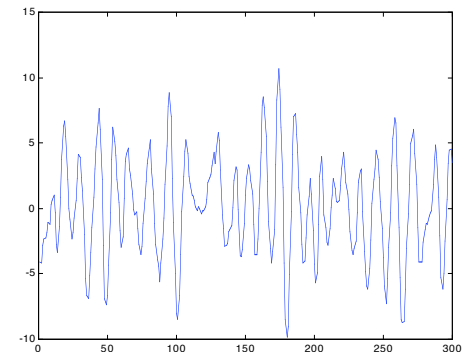


no correlation, no information



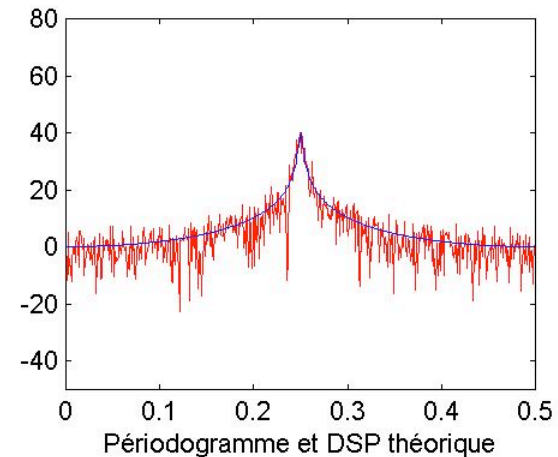
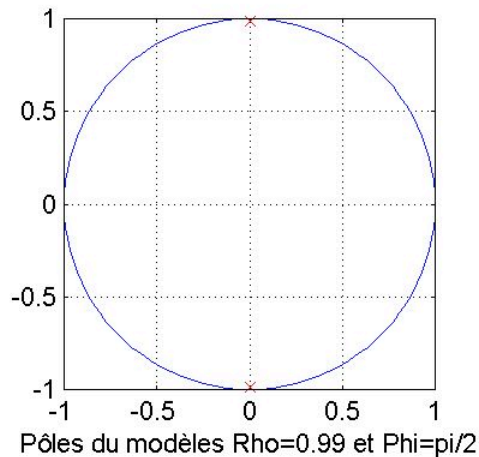
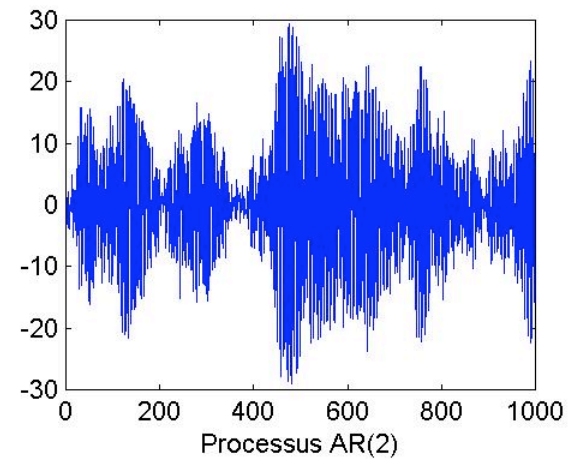
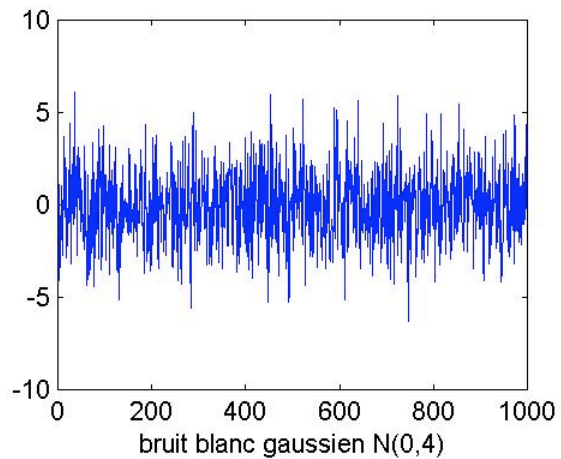
$$x(k) = - \sum_{i=1}^n a_i x(k-i) + \sum_{j=0}^m b_j n(k-j)$$

Model parameters: $(a_i, b_j \sigma_n^2)$?

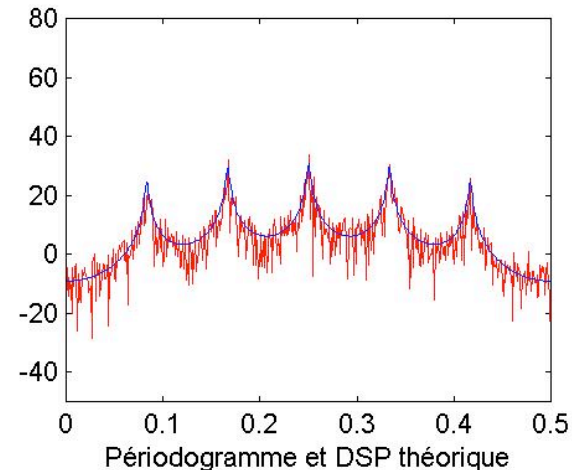
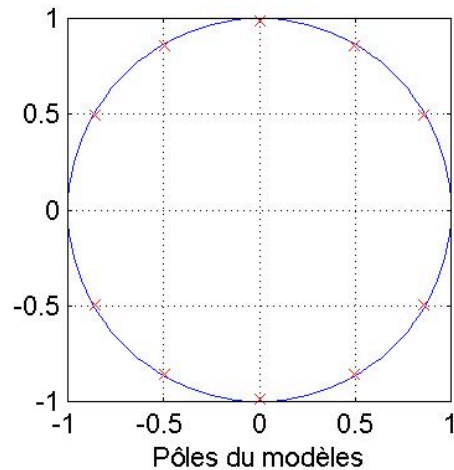
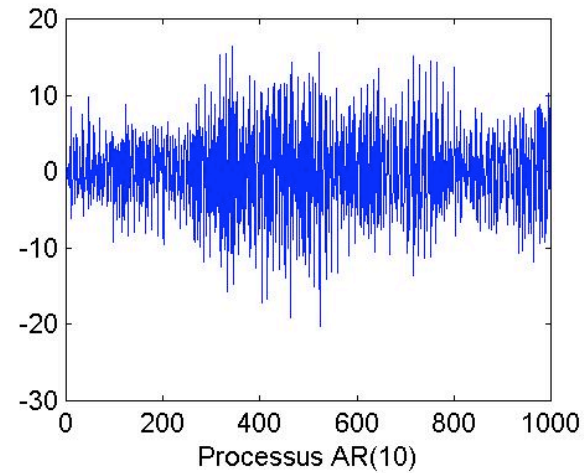
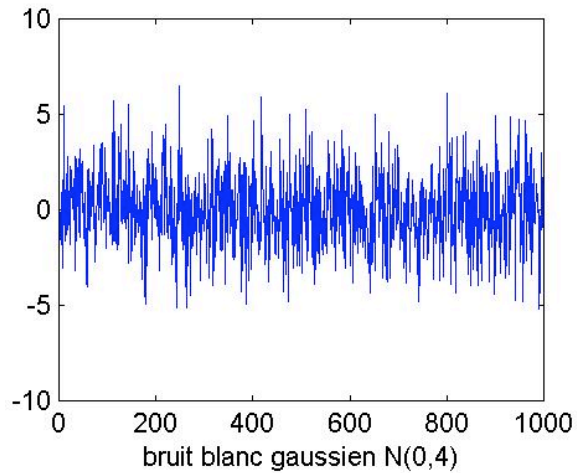


Correlation brought by the filter

Example of AR process

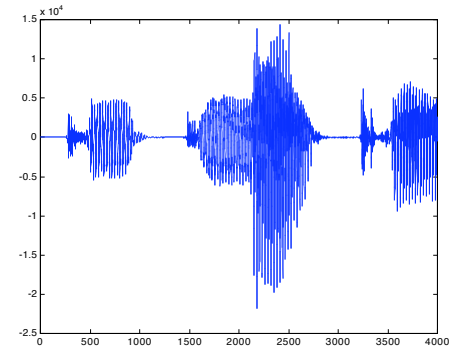
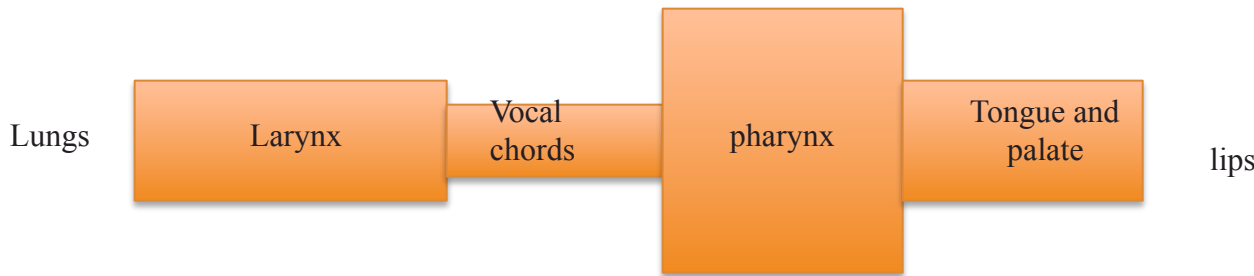


Example of AR process

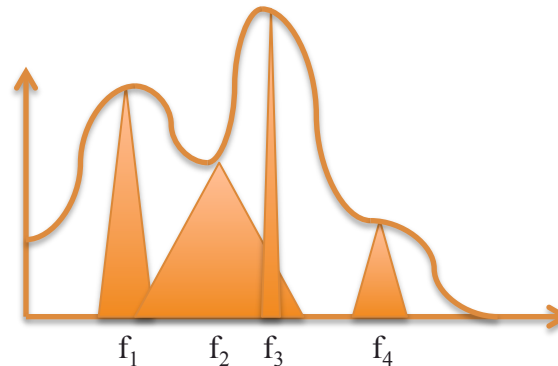
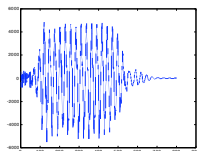
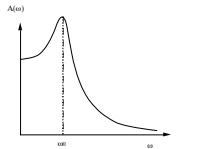
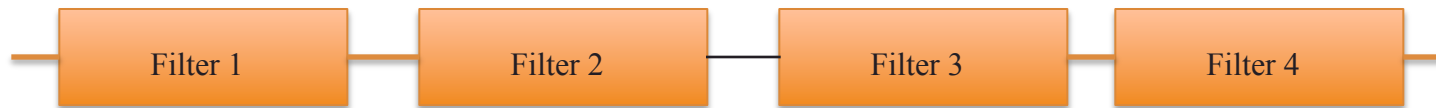


Regression models

■ Example : simplified physical speech model



■ Cascade of 4 resonator filters



Cumulative spectrum

Parametric model

- Autoregressive (AR) (all-poles filter)

$$x(k) = \sum_{i=1}^p a_i x(k-i) + n(k)$$

- Moving Average (MA) (all-zeros filter)

$$x(k) = \sum_{j=0}^q b_j n(k-j)$$

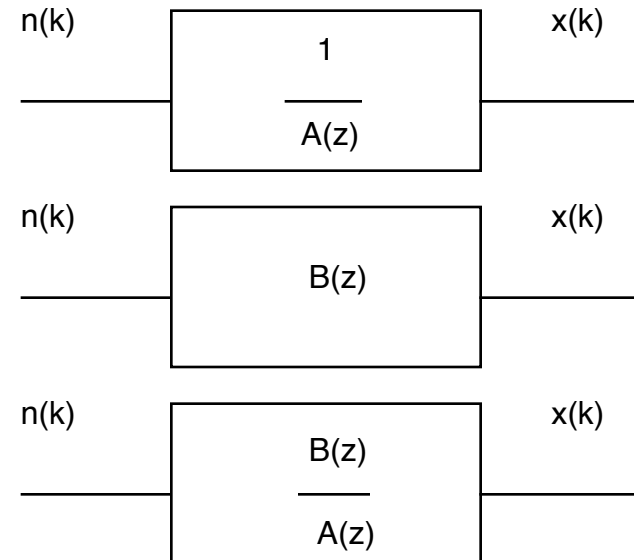
- ARMA (poles and zeros filter)

$$x(k) = \sum_{i=1}^p a_i x(k-i) + \sum_{j=0}^q b_j n(k-j)$$

- ARIMA (ARMA on $\delta x(k) = x(k) - x(k-1)$)

$$Z(x(k-i)) = z^{-i}Z(x(k))$$

- Z: delay operator
- n(k) : white noise with mean zero and correlation




$$\Gamma(0) = \sigma^2$$

$$\Gamma(k) = 0 \quad k \neq 0$$

Model estimation

- AR model estimation by linear prediction

p centered data observed from $(k-p)$ to $(k-1)$  linear prediction of the data value at time k

Predictive value

$$\hat{x}(k) = \sum_{i=1}^p a_i x(k-i)$$

Prediction error

$$e(k) = x(k) - \hat{x}(k) = x(k) - \sum_{i=1}^p a_i x(k-i)$$

Link with AR model

When is $e(k)$ a white noise?

Model estimation

- Least square

$$\min_{a_i} \sum e^2(k)$$

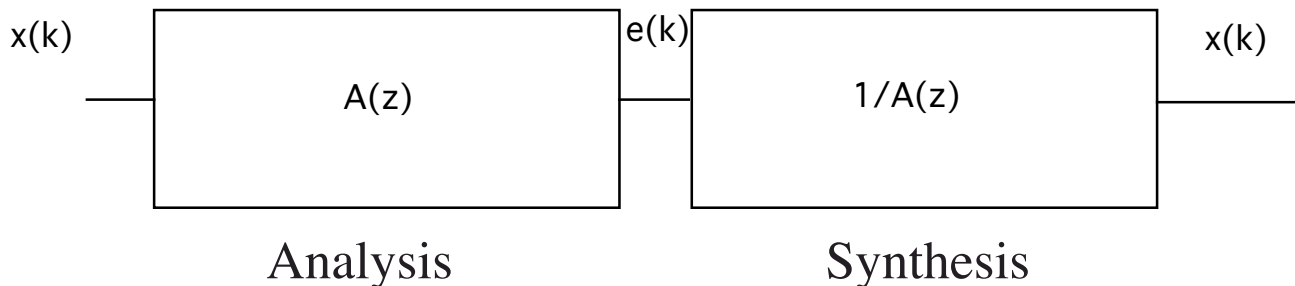
$$\text{grad}(\sum e^2(k)) = 0$$

$$\sum e(k) \frac{\partial e(k)}{\partial a_i} = 0 \quad i = 1, \dots, p$$

$$\sum e(k)x(k-i) = 0 \quad i = 1, \dots, p$$

- ➡ no correlation between error and data
- ➡ decorrelation of $x(k)$ and $e(k)$ tends to a white noise

- Inverse filtering



Model estimation

- Yule-Walker equations

$$\sum e(k)x(k-i) = 0 \quad i = 1, \dots, p$$

$$\sum_{k=1}^n (x(k) - \sum_{j=1}^p a_j x(k-j))x(k-i) = 0 \quad i = 1, \dots, p$$

$$\sum_{j=1}^p a_j \hat{\Gamma}_x(i-j) = \hat{\Gamma}_x(i) \quad i = 1, \dots, p$$

- Linear system of equations

$$\begin{bmatrix} \Gamma(0) & \dots & \dots & \Gamma(p-1) \\ \cdot & \cdot & & \cdot \\ \cdot & & \cdot & \cdot \\ \Gamma(p-1) & \dots & \dots & \Gamma(0) \end{bmatrix} \begin{bmatrix} a_1 \\ \cdot \\ \cdot \\ a_p \end{bmatrix} = \begin{bmatrix} \Gamma(1) \\ \cdot \\ \cdot \\ \Gamma(p) \end{bmatrix} \quad \longrightarrow \quad \begin{bmatrix} a_1 \\ \cdot \\ \cdot \\ a_p \end{bmatrix} = \begin{bmatrix} \Gamma(0) & \dots & \dots & \Gamma(p-1) \\ \cdot & \cdot & & \cdot \\ \cdot & & \cdot & \cdot \\ \Gamma(p-1) & \dots & \dots & \Gamma(0) \end{bmatrix}^{-1} \begin{bmatrix} \Gamma(1) \\ \cdot \\ \cdot \\ \Gamma(p) \end{bmatrix}$$

$$\sigma_e^2 = \sigma_x^2 - \sum_{i=1}^p a_i \Gamma(i)$$

Parametric model

■ Estimation of AR coefficients

- Estimation of covariance functions

$$\hat{\Gamma}(k) = \frac{1}{n} \sum_{l=0}^{n-1} x(l)x(l+k)$$

- Computation of coefficients and error variance by Yule-Walker equations

■ Associated power spectral estimate

$$S_x(f) = \frac{\sigma_e^2}{\left| \sum_{i=0}^p a_i e^{-j2\pi i f} \right|^2}$$

Example of parametric models

■ AR 1

$$x(k) = a_1 x(k-1) + n(k)$$

$$a_1 = \frac{\Gamma(1)}{\Gamma(0)} \simeq \frac{\sum_{i=0}^{n-1} x(i)x(i+1)}{\sum_{i=0}^{n-1} x(i)^2}$$

$$\sigma_n^2 = \sigma_x^2 - a_1 \Gamma(1) = \sigma_x^2 - \frac{\Gamma(1)^2}{\sigma_x^2}$$

■ AR 2

$$x(k) = a_1 x(k-1) + a_2 x(k-2) + n(k)$$

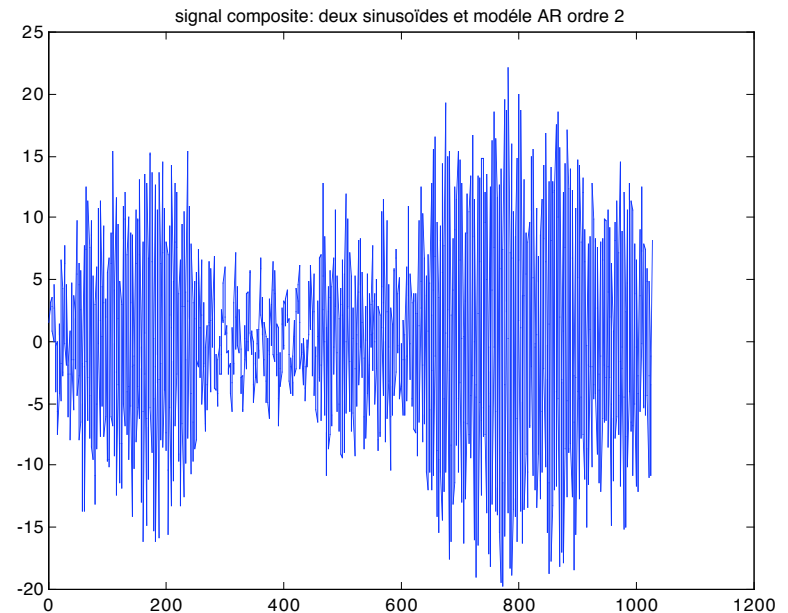
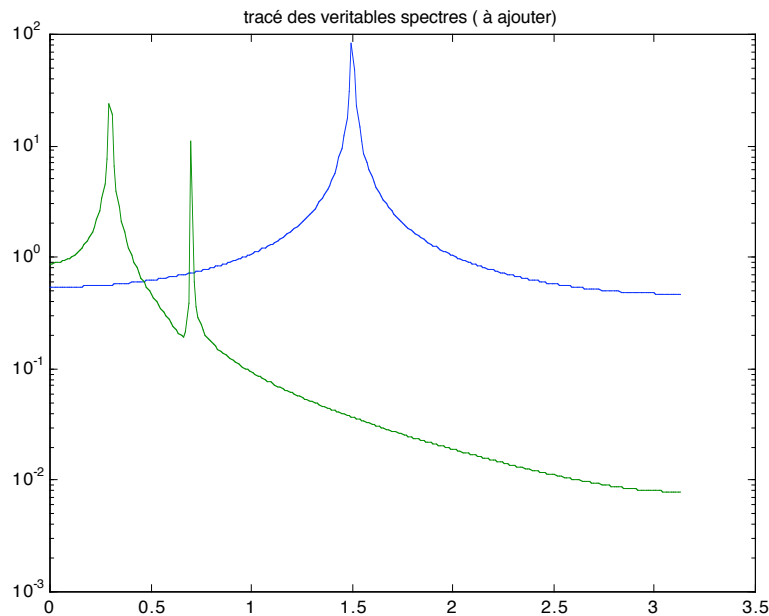
$$a_1 = \frac{\Gamma(1)\Gamma(0) - \Gamma(1)\Gamma(2)}{\Gamma(0)^2 - \Gamma(1)^2}$$

$$a_2 = \frac{\Gamma(1)^2 - \Gamma(0)\Gamma(2)}{\Gamma(0)^2 - \Gamma(1)^2}$$

$$\sigma_n^2 = \sigma_x^2 - a_1 \Gamma(1) - a_2 \Gamma(2)$$

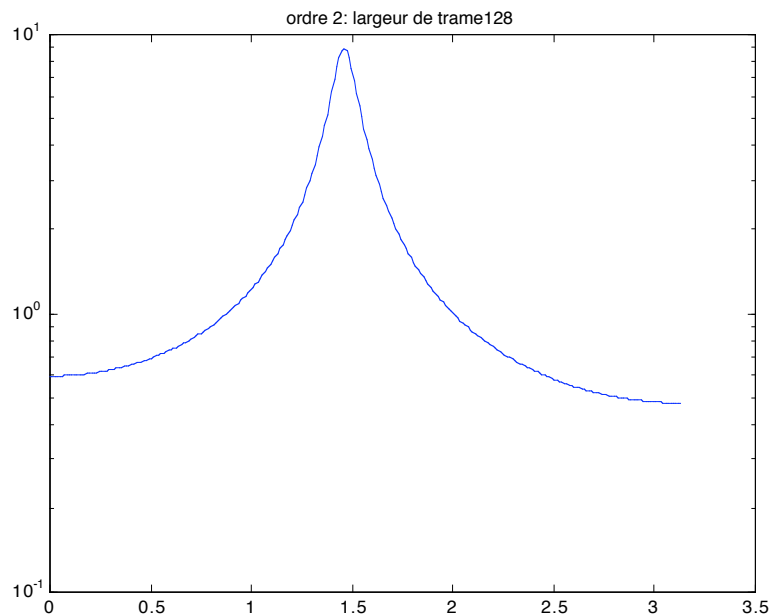
Parametric model

- Example
 - Complex signal as a sum of two sine waves and a stationary stochastic process with a damped correlation

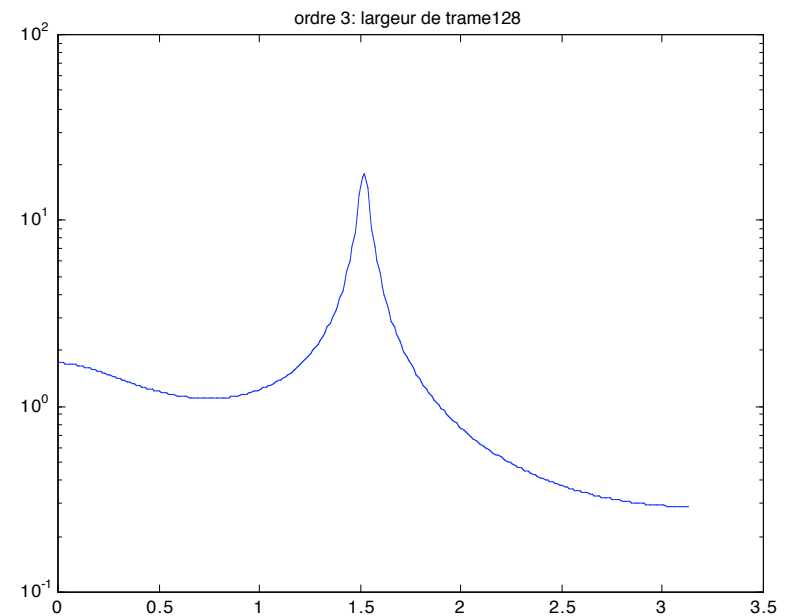


Parametric model

- AR model



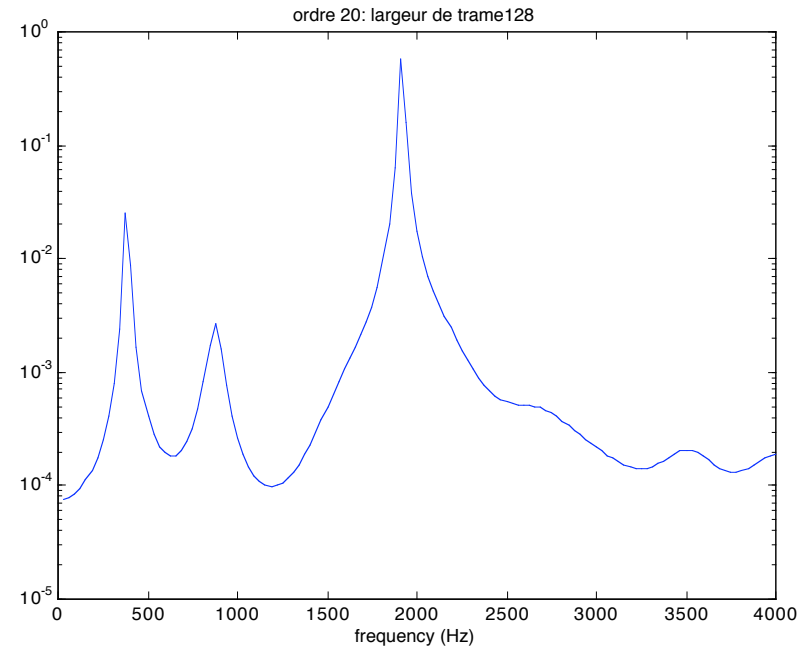
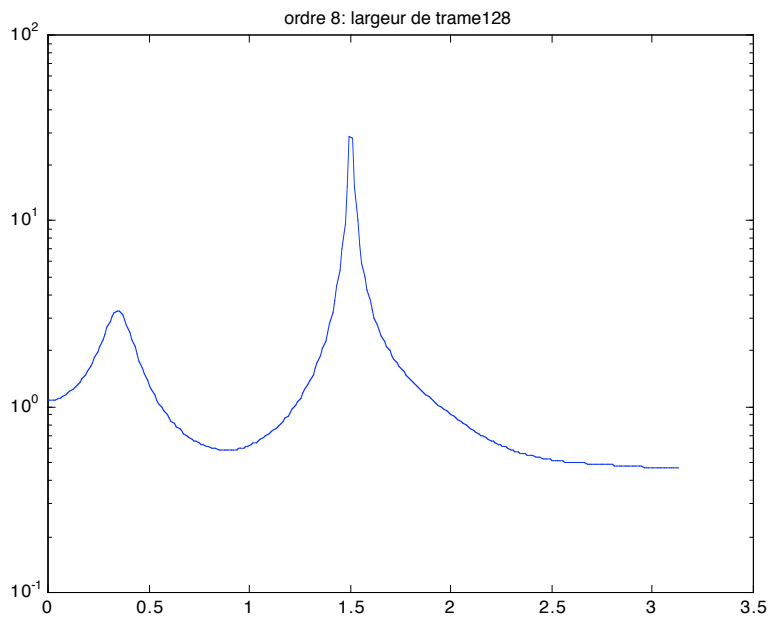
P=2, n=128



P=3, n=128

Parametric model

- AR model



P= 8,n=128

P=20,n=128

Model identification (order estimation)

- Minimization of a function of p given by information-theoretic approach
 - Rissanen or Bayesian Information Criterion (BIC) for gaussian data

$$RIS(p) = \ln(\sigma_e^2(p)) + p \frac{\ln(n)}{n}$$

- Akaike Information criterion (AIC)

$$AIC(p) = \ln(\sigma_e^2(p)) + \frac{2p}{n}$$

Expansion on orthogonal basis functions

■ Advantages

- More compact representation of the signal energy
- Convenient for characterizing dominant pattern of variability

■ Time series

$$x(k) = \sum_{i=1}^n c_i f_i(k)$$

$$c_i = \sum_{k=1}^n x(k) f_i^*(k)$$

- where

$$\sum_k f_i(k) f_j^*(k) = \delta_{ij}$$

$$\delta_{ij} = 1 \quad i = j, \delta_{ij} = 0 \quad i \neq j$$

- Complex exponentials are a special case and define a set of orthonormal functions

$$f_i(k) = \frac{1}{\sqrt{n}} e^{j \frac{2\pi i k}{n}}$$

- Fourier transform is also an expansion on orthonormal basis

Empirical Orthogonal Functions (EOF)

- Coefficients must be uncorrelated random variables

$$E[c_i c_j^*] = \lambda_i \delta_{ij}$$

- Solution

$$E[c_i c_j^*] = \sum_k \sum_l E[x(k)x(l)] f_i(k) f_j^*(l) = \lambda_i \delta_{ij}$$

- as $\sum_k f_i(k) f_j^*(k) = \delta_{ij}$

$$\sum_k E[x(k)x(l)] f_i(k) = \lambda_i f_i(l)$$

Empirical Orthogonal Functions (EOF)

- Stationary case

$$\Gamma(k - l) = E[x(k)x(l)]$$

$$\sum_{k=1}^n \Gamma(k - l) f_i(k) = \lambda_i f_i(l) \quad l = 1, \dots, n$$

$$\begin{bmatrix} \Gamma(0) & \cdots & \Gamma(n-1) \\ \vdots & \ddots & \vdots \\ \Gamma(n-1) & \cdots & \Gamma(0) \end{bmatrix} \begin{bmatrix} f_i(1) \\ \vdots \\ f_i(n) \end{bmatrix} = \lambda_i \begin{bmatrix} f_i(1) \\ \vdots \\ f_i(n) \end{bmatrix}$$

➔ Solution : Eigenvectors and eigenvalues of the covariance matrix

$$\Gamma\Phi = \Lambda\Phi$$

$$\Lambda = \begin{bmatrix} \lambda_1 & \cdots & 0 \\ \vdots & \ddots & \vdots \\ 0 & \cdots & \lambda_n \end{bmatrix}$$

$$\Phi = [\Phi_1 \quad \cdots \quad \Phi_n]$$

$$\Phi_i = [f_1 \quad \cdots \quad f_n]^T$$

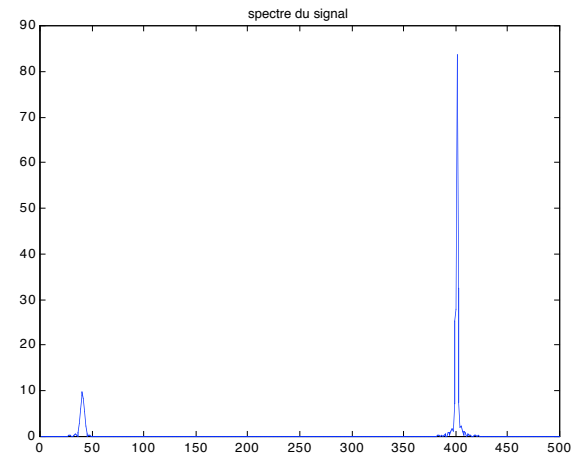
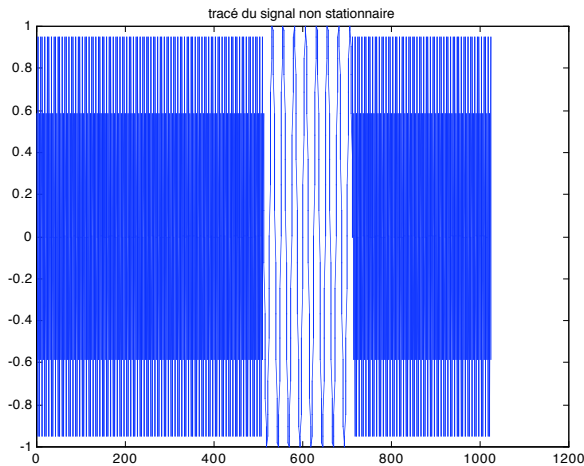
Other names for this method

- **Principal Components Analysis (PCA)**
- **Kahrunen-Loeve decomposition (continuous signals)**
- **Diagonalization of a covariance matrix**
 - Eigenvectors give the principal directions of vector clusters
 - Eigenvalues give the variances along each principal direction
- **Singular Spectrum Analysis**

Time-frequency analysis

■ Problems with Fourier analysis

- No spatial information

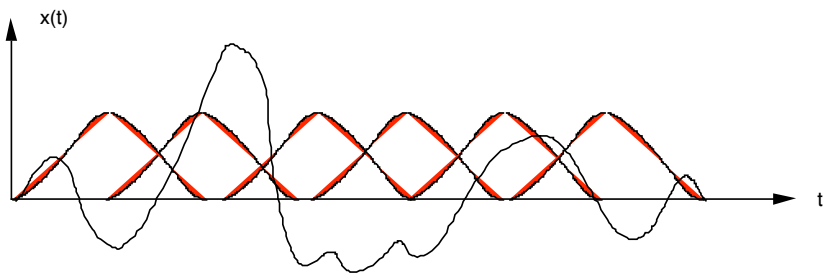


- Not a realistic model (destructive interferences)

$$x(t) = \int_{-\infty}^{+\infty} X(f) e^{j2\pi ft} df = 0 \quad t \notin [0, T]$$

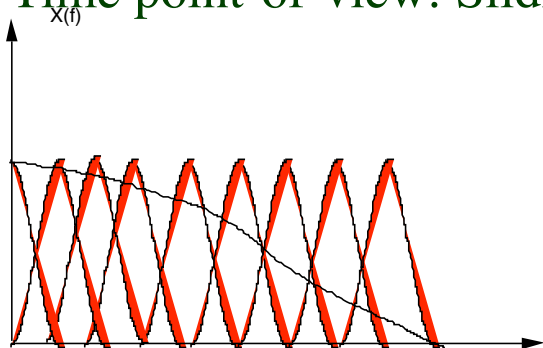
Time-frequency analysis

■ Short-Time Fourier transform



analog to a new transform
with function $g(s)$

Time point-of-view: Sliding window



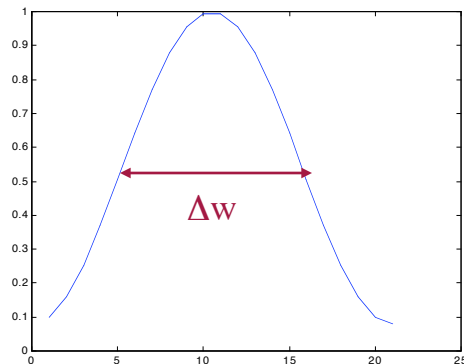
$$g_t(s) = w(s - t)e^{j2\pi fs}$$

$$X(t, f) = \int_{-\infty}^{\infty} x(s)w^*(s - t)e^{-j2\pi fs} ds$$

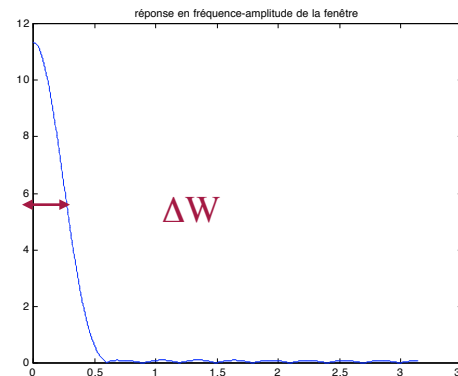
$$X(t, f) = \int_{-\infty}^{\infty} x(s)g_t^*(s, f) ds$$

Frequency point-of-view: Bandpass filter bank

■ Uncertainty principle



$$\Delta w \cdot \Delta W \geq \frac{1}{4\pi}$$



Time window

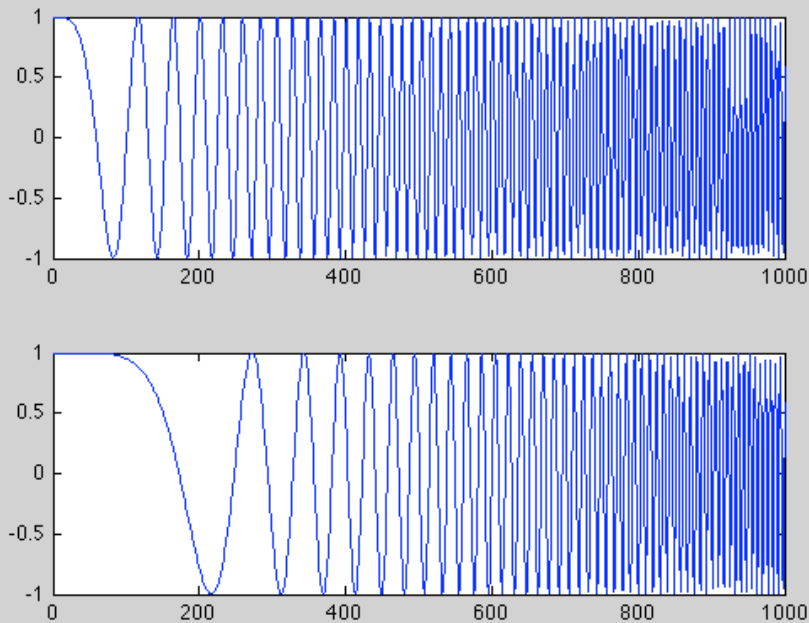
Window Spectrum

- Good time resolution, bad spectrum resolution
- Good spectrum resolution, bad time resolution

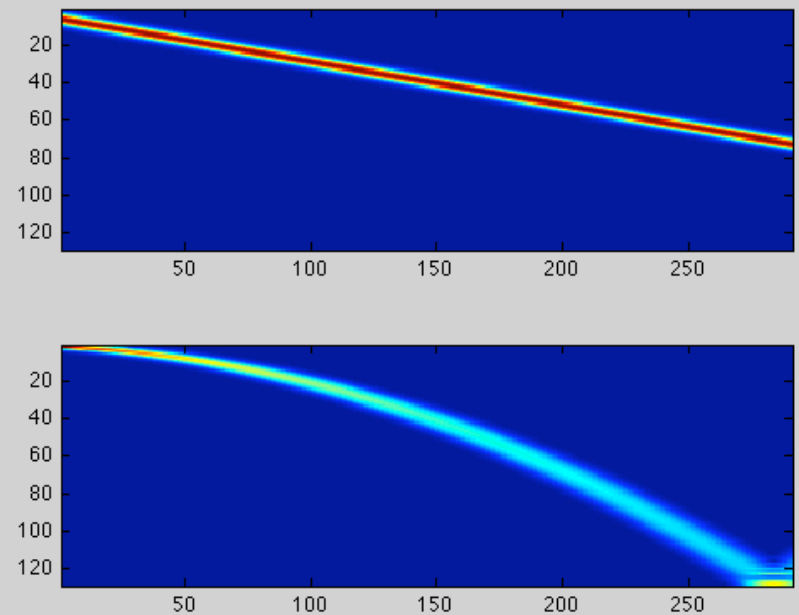
Time frequency analysis

■ Spectrogram (spectrum function of time)

Linear Chirp (up) and quadratic chirp (bottom)



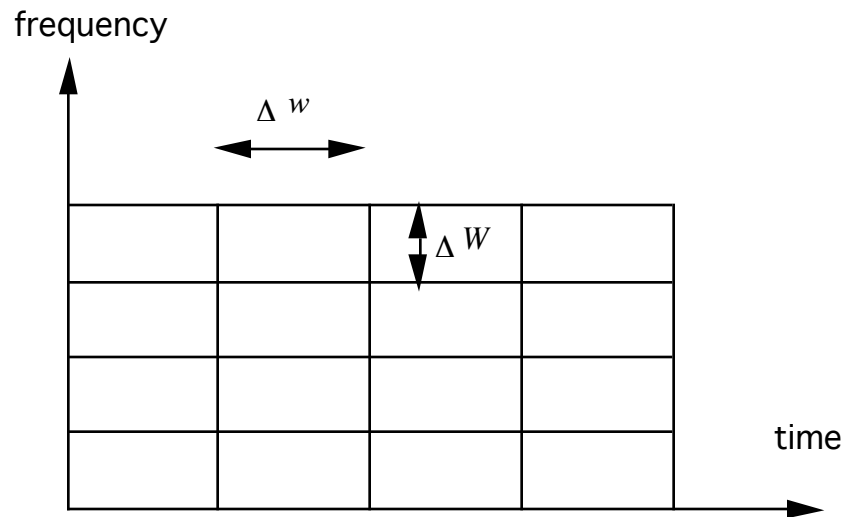
Signal



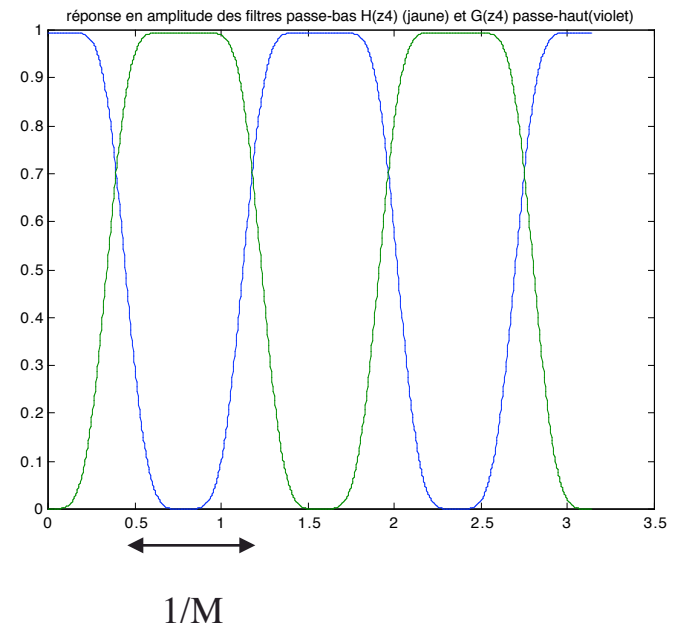
Spectrogram

Time frequency analysis analysis

■ Modulated filter bank



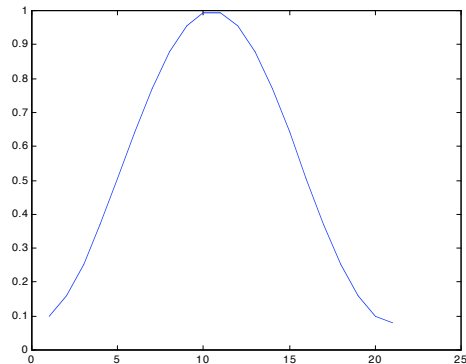
Time-frequency plane tiling



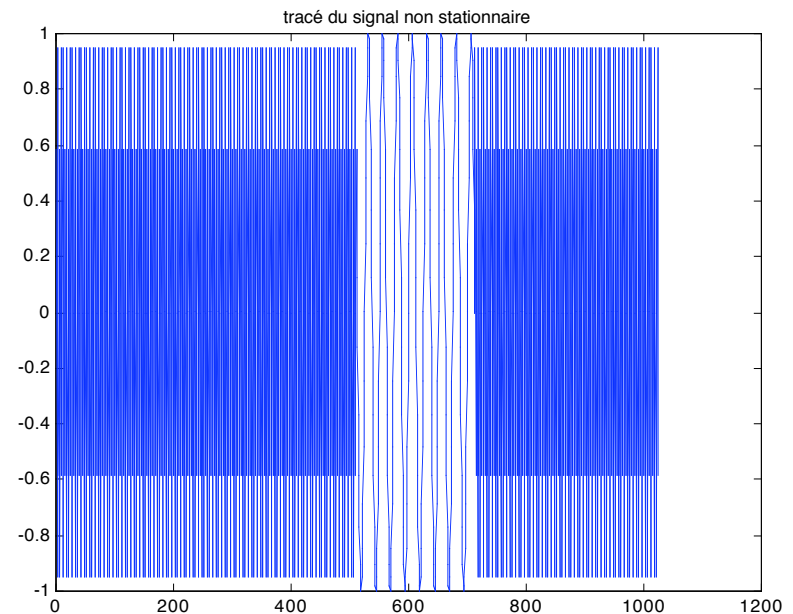
Frequency point-of-view:
Modulated filterbank

Time-frequency analysis

■ Short-time Fourier Transform

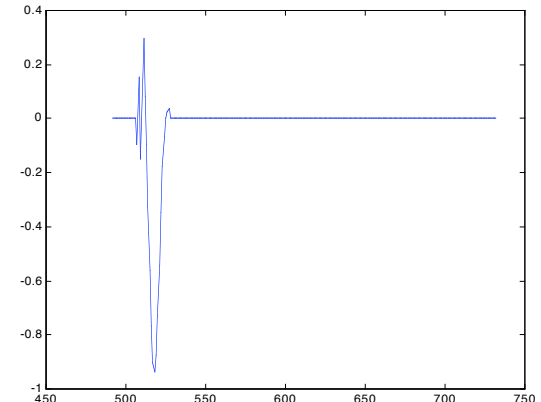
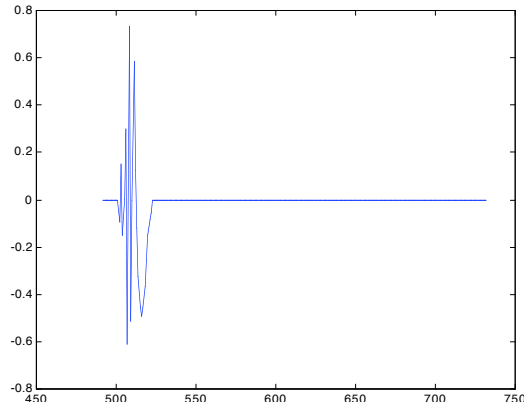
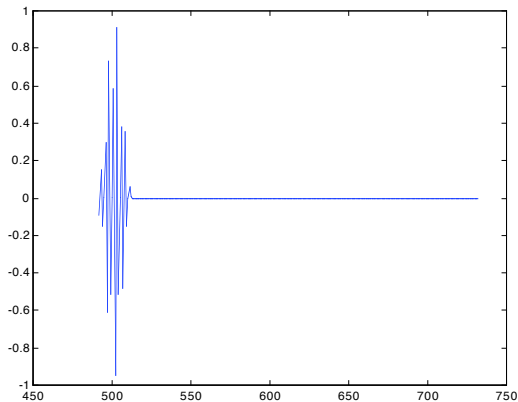


Hamming window

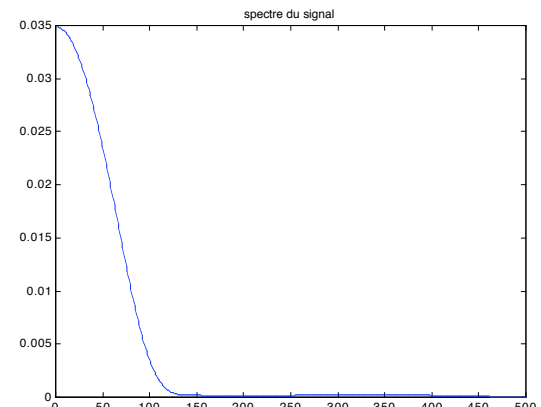
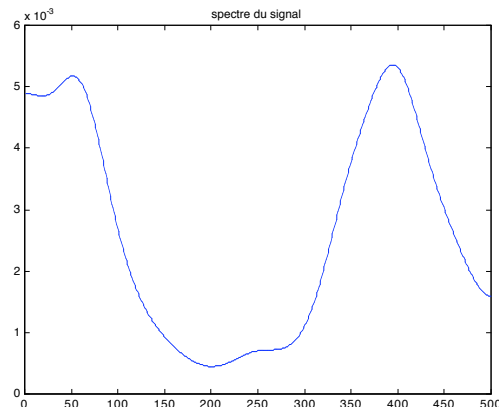
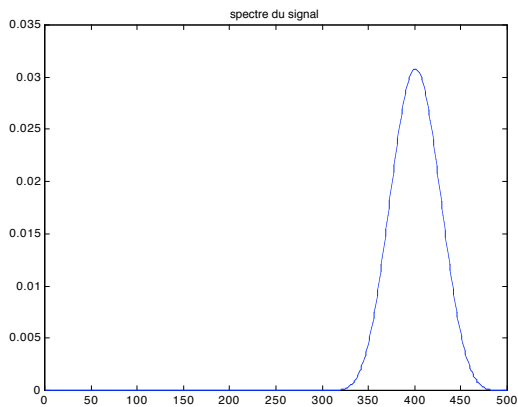


Non stationary signal

Time frequency analysis



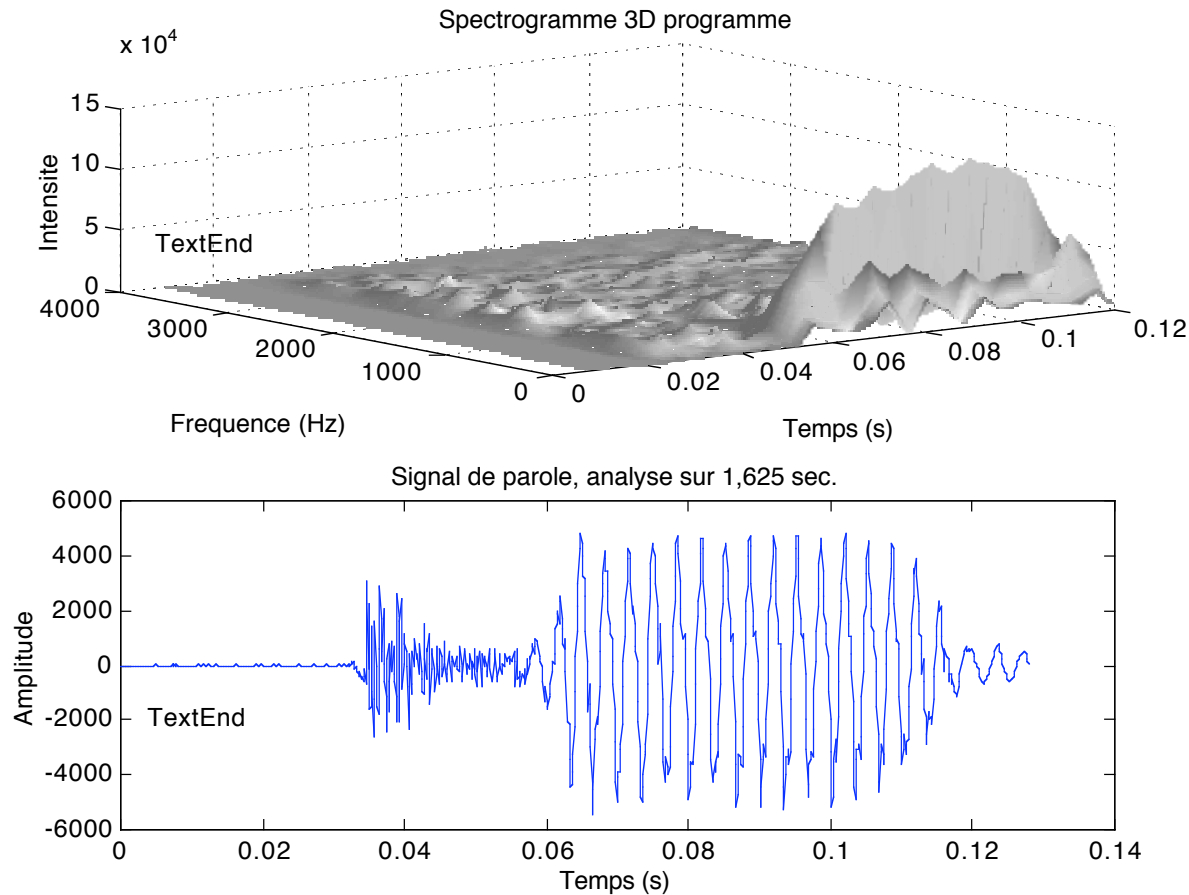
Signal



Spectrum

Time frequency analysis

■ Speech signal: spectrogram



Time-Scale analysis

■ Statement

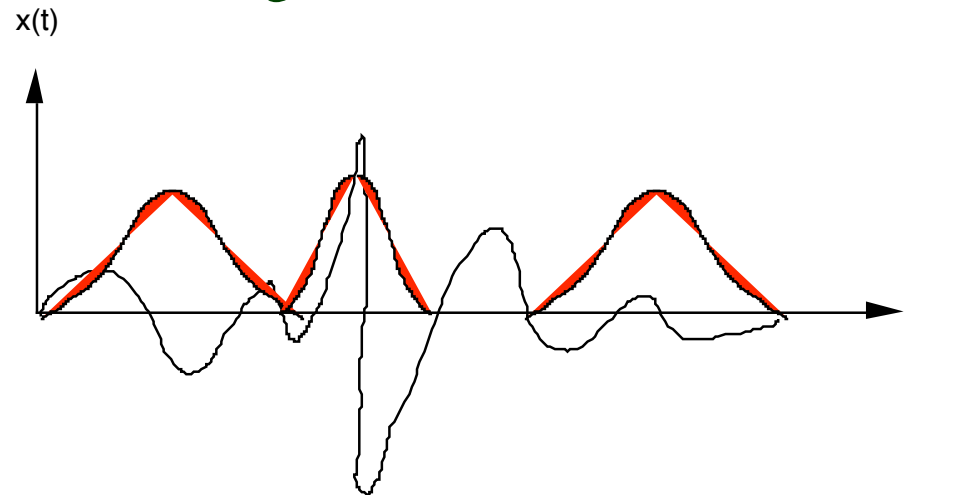
- The information, which is sought in a signal, varies.

Transients, bursts, fast changes : fine time resolution

Periodic signals, long term evolution : fine frequency resolution



Sliding window with different time support



- Large window for long term evolution , Short window for transients

Continuous wavelet transform

$\psi(t)$ is a normalized window called « wavelet » with zero mean

$$\int \psi(t) dt = 0 \qquad \int |\psi(t)|^2 dt = 1$$

New transformation over dilated and translated wavelets

$$\psi_{a,b}(t) = \frac{1}{\sqrt{a}} \psi\left(\frac{t-b}{a}\right) \qquad b \in \mathbb{R}, a > 0$$

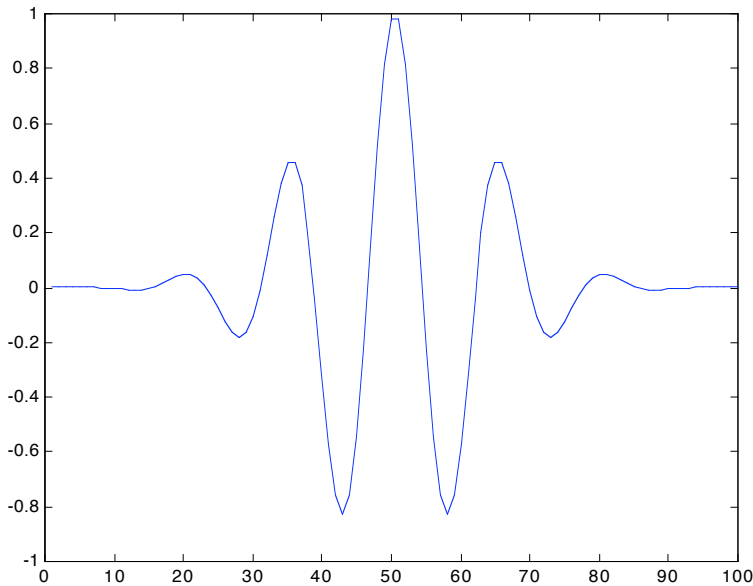
$$C_{a,b} = \int_{-\infty}^{\infty} x(t) \psi_{a,b}^*(t) dt$$

a: scale factor

b : translation factor

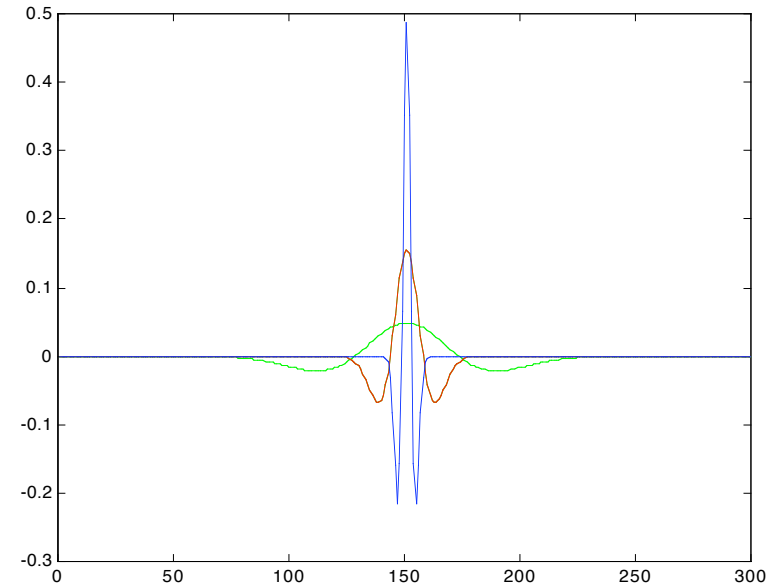
Different wavelets

■ Morlet wavelet



$$\psi(t) = e^{-t^2/2} \cos(5t)$$

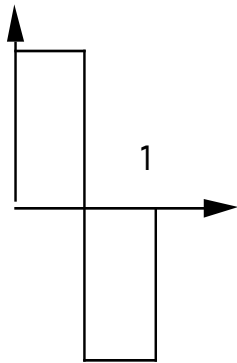
◆ Mexican hat



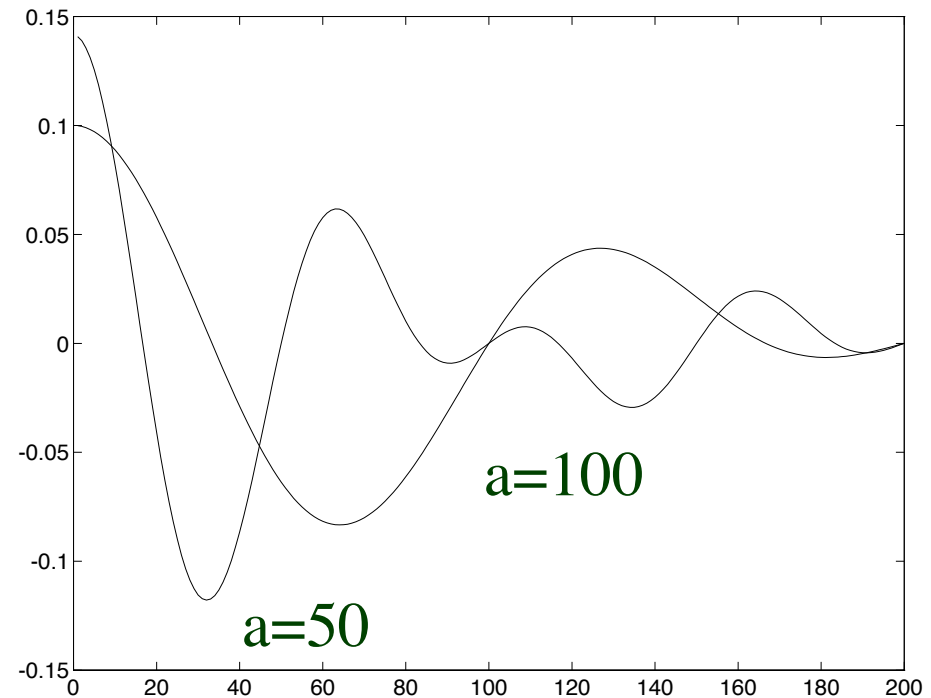
- ◆ a=0.1 ◆ Second derivative of gaussian
 - ◆ a=0.01
 - ◆ a=0.001
- $$\psi(t) = \left(\frac{2}{\sqrt{3}\sqrt{\pi}} \right) (1-t^2) e^{-t^2/2}$$

Different wavelets

■ Haar wavelet



◆ Shannon wavelet



Wavelet transform

■ Inverse continuous wavelet transform

$$x(t) = \frac{1}{C_\psi} \iint C_{a,b} \psi_{a,b}(t) \frac{da db}{a^2}$$

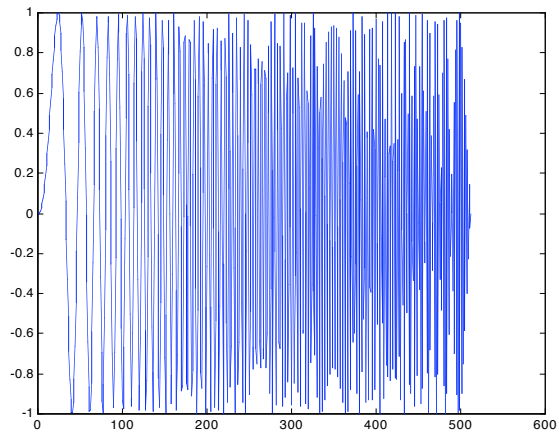
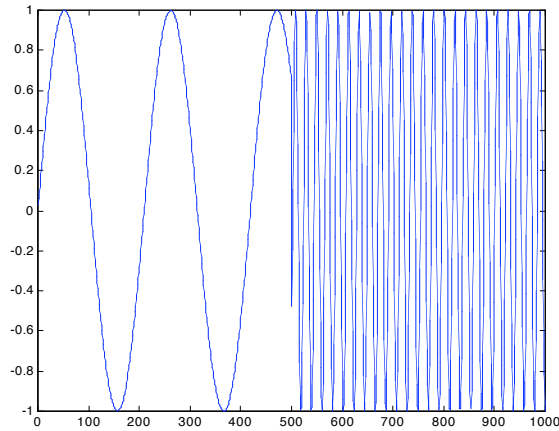
Admissibility condition: analysis and reconstruction without error

$$C_\psi = \int \frac{|\Psi(\omega)|^2}{|\omega|} d\omega < \infty \quad \longrightarrow \quad \Psi(0) = \int \psi(t) dt = 0$$

■ A wavelet is a symmetric time window with null mean

Example

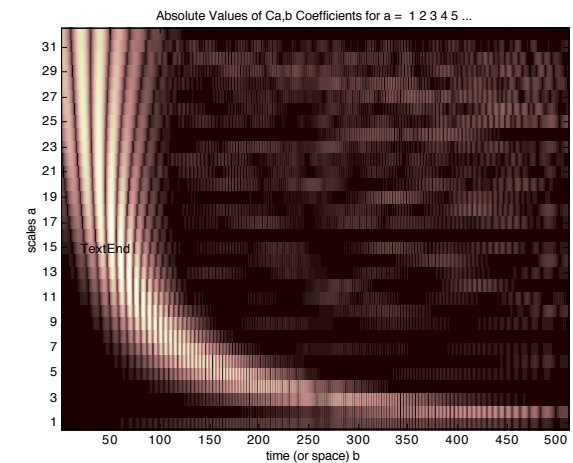
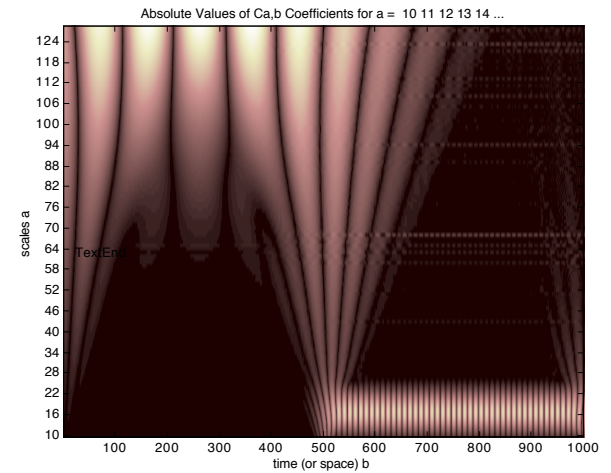
Signal



Breakdown
of sinusoids

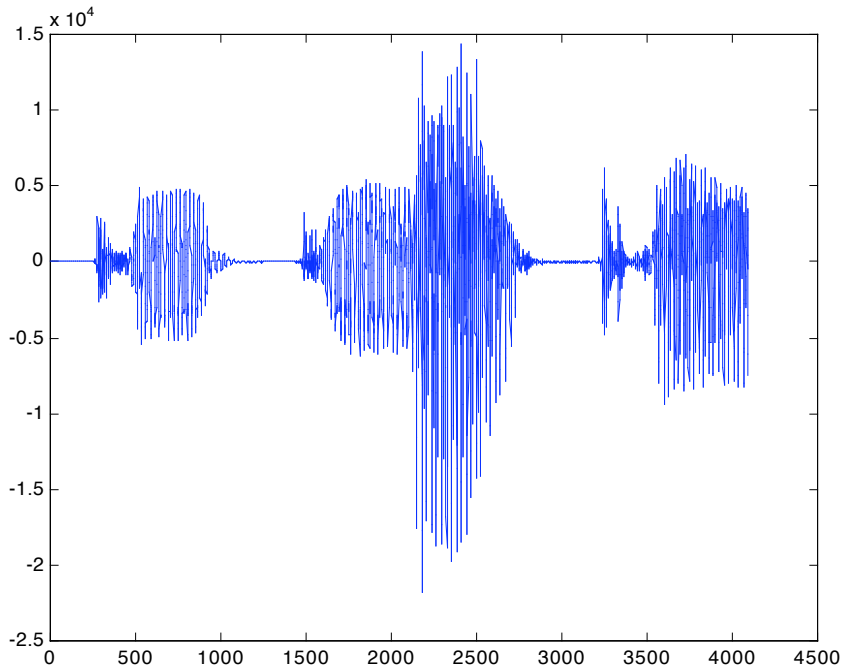
Linear chirp

Scalogram

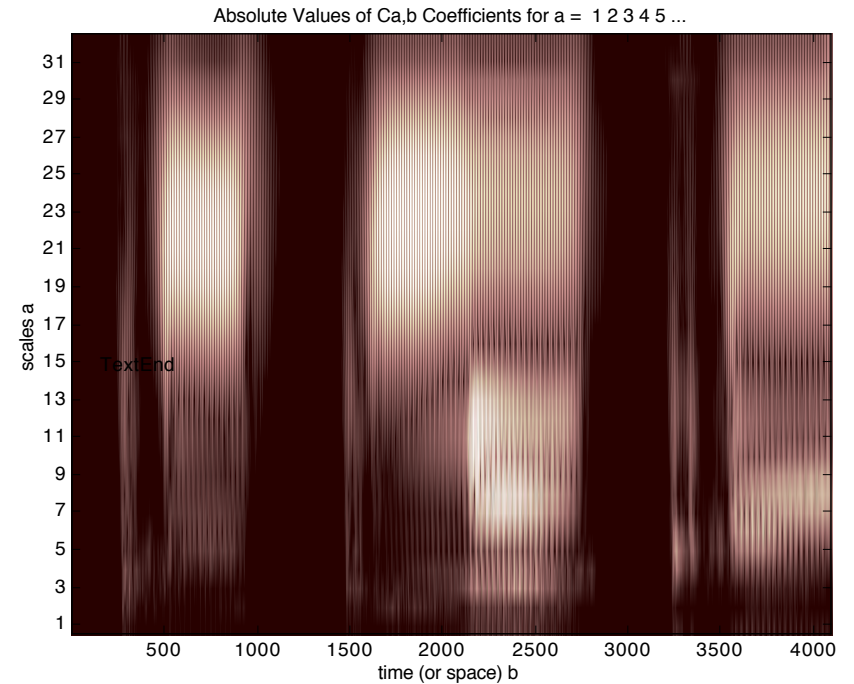


Example

Speech Signal



Scalogram (Morlet) $a=1:32$



$a=1$: high frequencies, short scales
 $a=32$: low frequencies, large scales

Expansion on orthonormal basis

- Special choice of sampling in time and scale leads to orthonormal wavelets

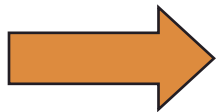
$$\int \psi_{j,k}(t) \psi_{n,m}(t) dt = \begin{cases} 1 & j = n, k = m \\ 0 & \text{ailleurs} \end{cases}$$

- ◆ This is the case of a dyadic sampling in scale (power of 2)

- ◆ $a=2^j$

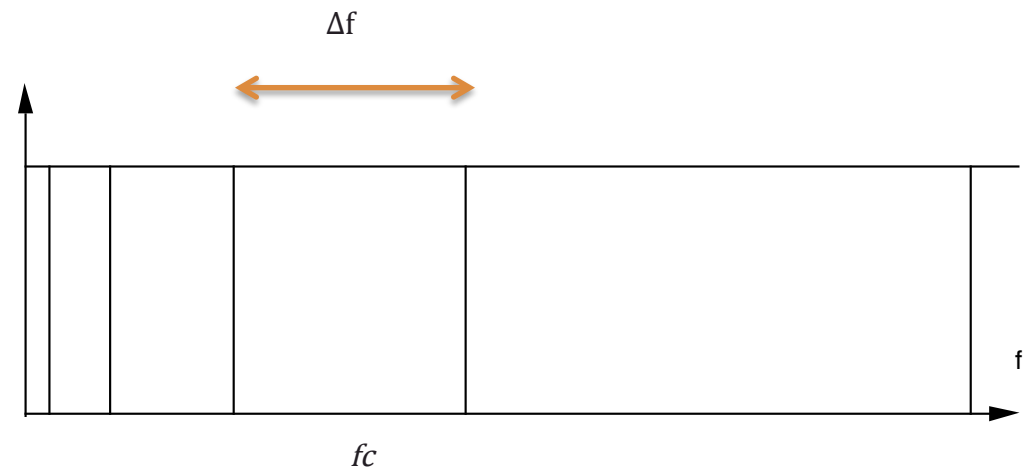
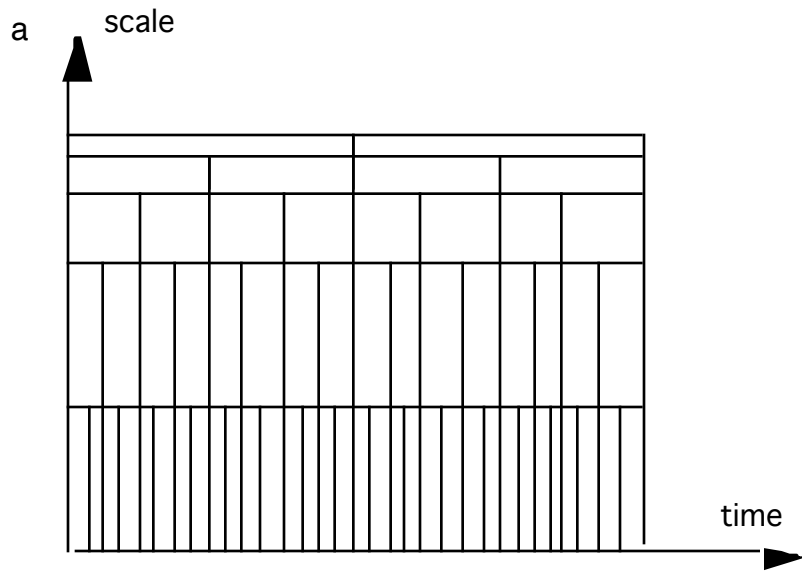
- ◆ $b=a.k$

$$\Psi_{j,k}(t) = 2^{-j/2} \Psi(2^{-j}t - k)$$



Non redundant representation of signal

Time-scale analysis

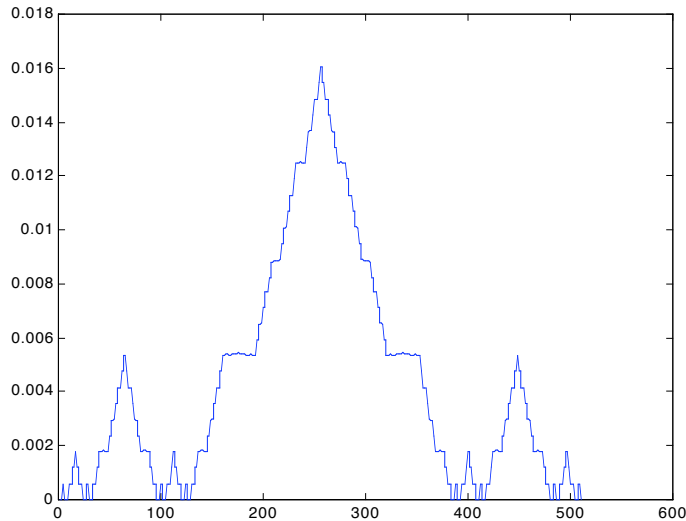


◆ Time-scale tiling

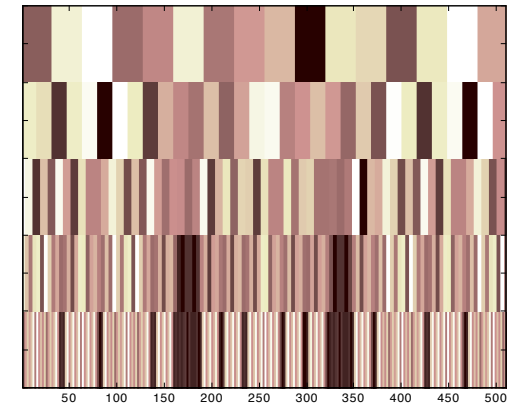
◆ Dyadic filter bank: $\frac{\Delta f}{f_c} = \text{constant}$

Time-scale analysis

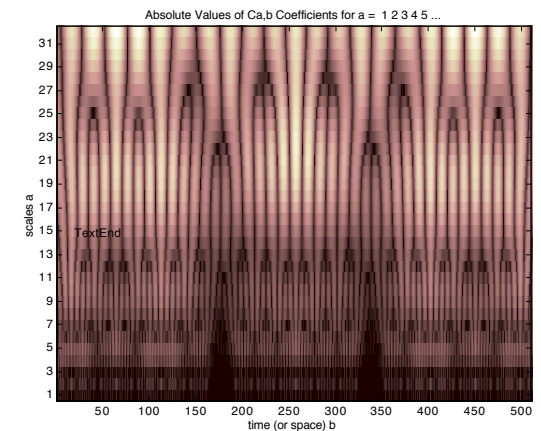
Fractal signal : Von Koch curve (flake)



Discrete
Scalogram
5 levels,
Daubechies 4



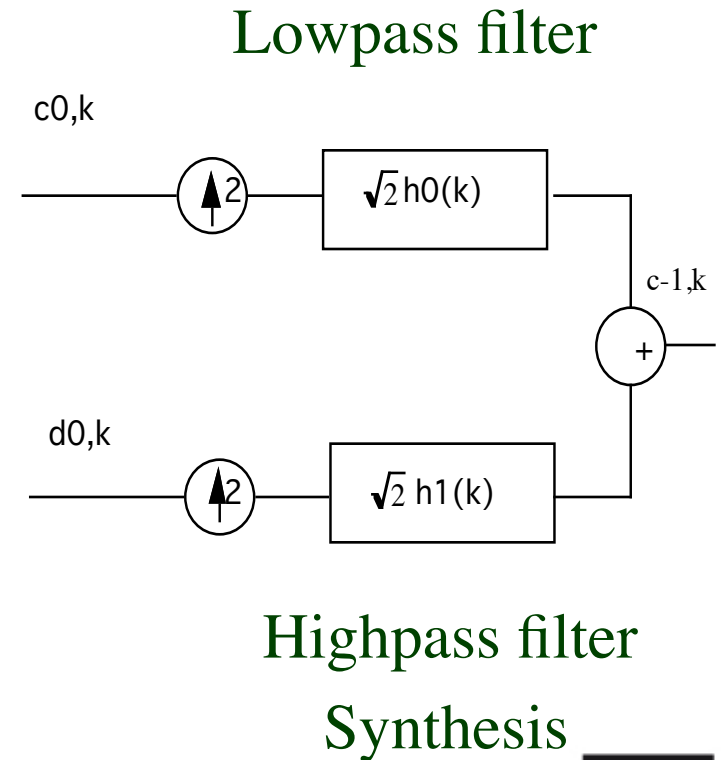
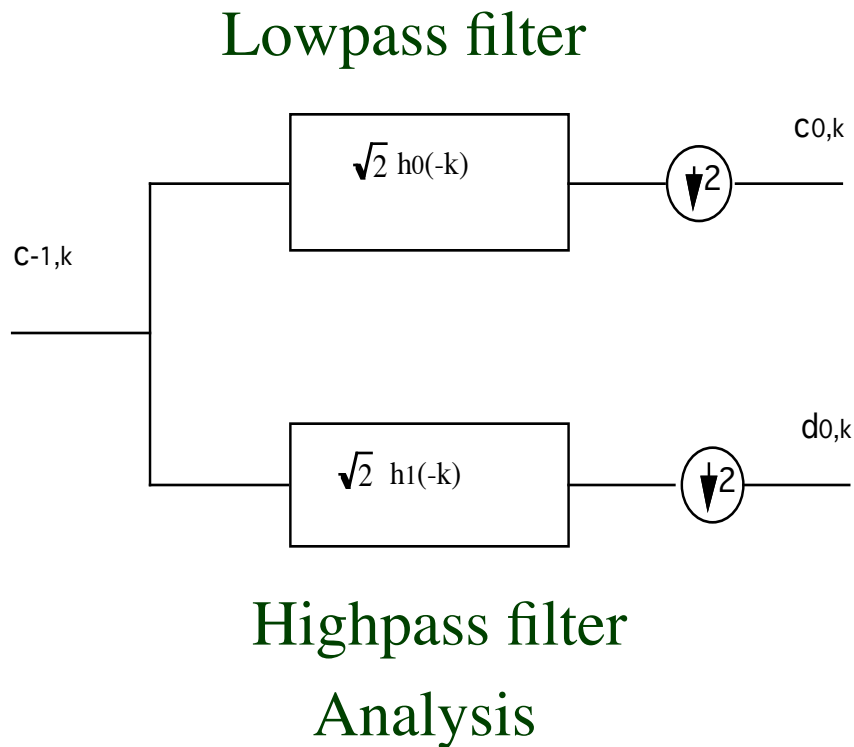
Continuous
Scalogram,
Morlet



Discrete wavelet implementation

■ Decomposition and reconstruction of signal between scale 0 and -1 by Mallat's algorithm

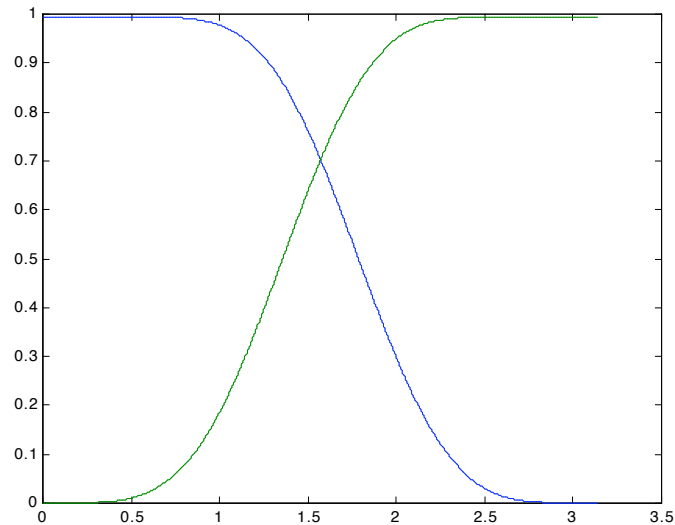
- No need for wavelet function, only filter coefficients



Discrete wavelet implementation

■ Conjugate mirror filters (Daubechies filters)

The lowpass and highpass filters have symmetrical transfer function for $\omega=\pi/2$



Condition of orthogonality for reconstruction without errors

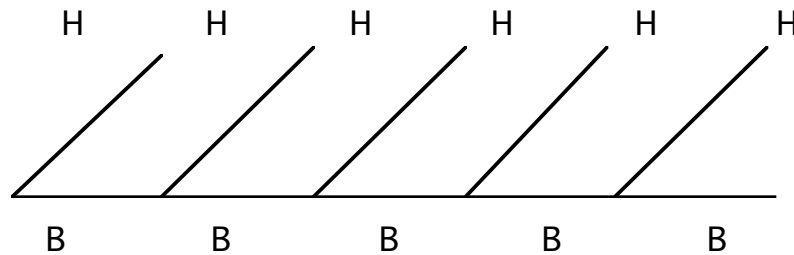


$$\left|H_0(e^{j\omega})\right|^2 + \left|H_1(e^{j\omega})\right|^2 = 1$$

Discrete wavelet analysis by Mallat's algorithm

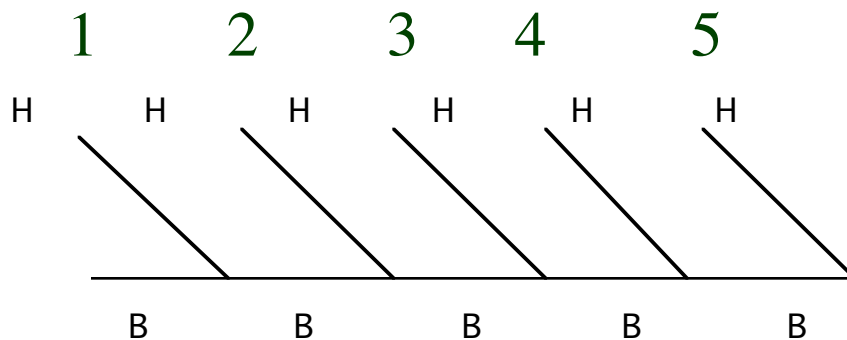
$$c_{0,k} = s(k)$$

B: lowpass filter (approximation)
H: highpass filter (details)



Analysis

Level 0

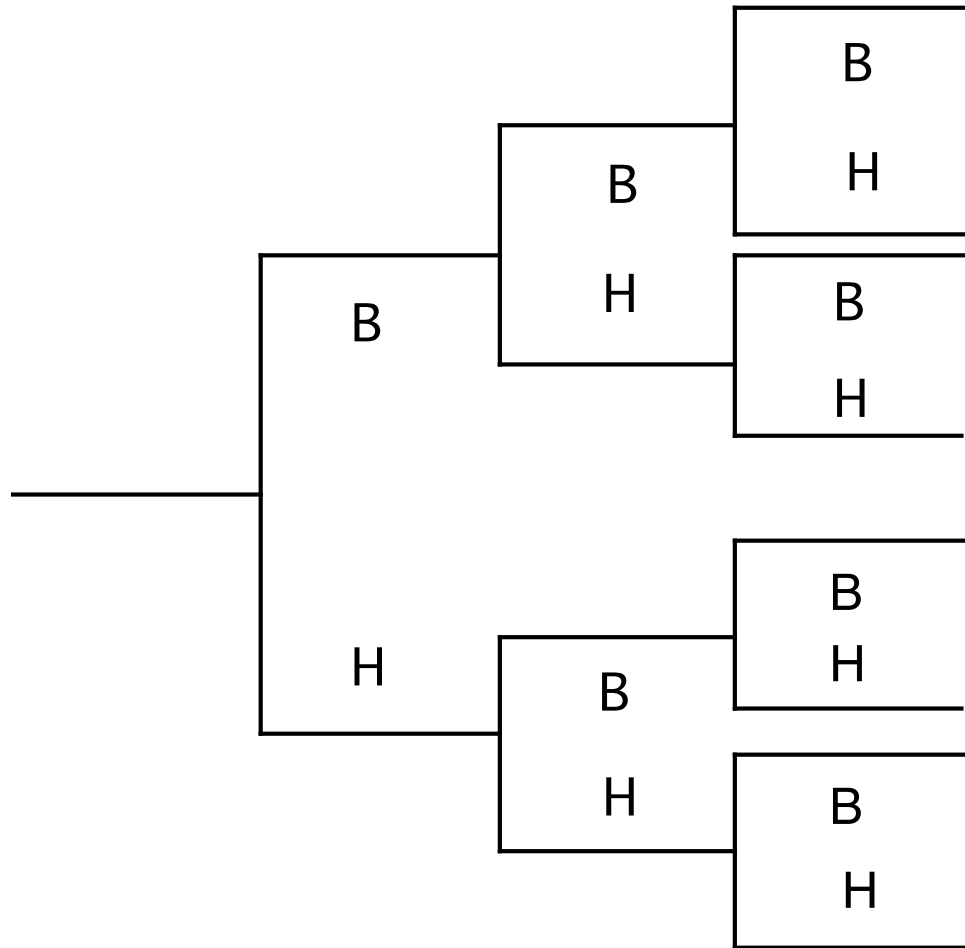


Synthesis

Level

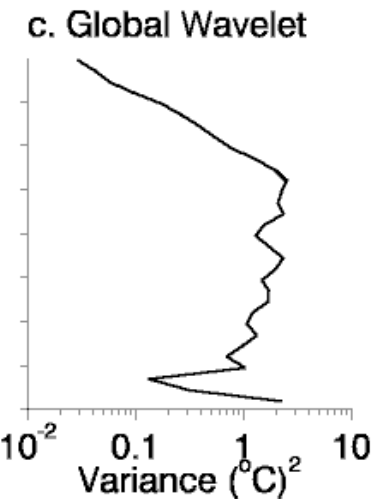
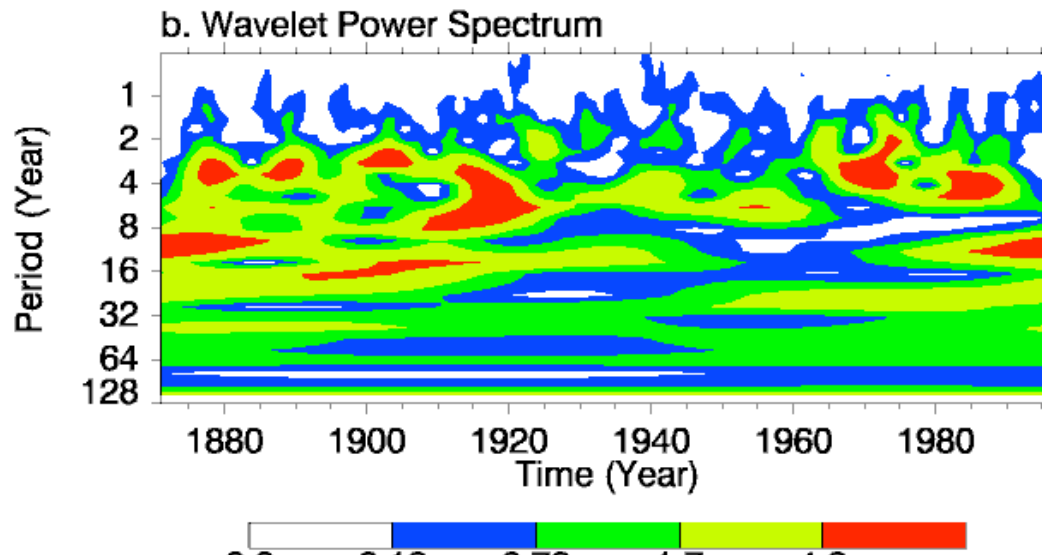
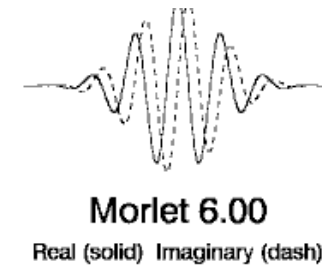
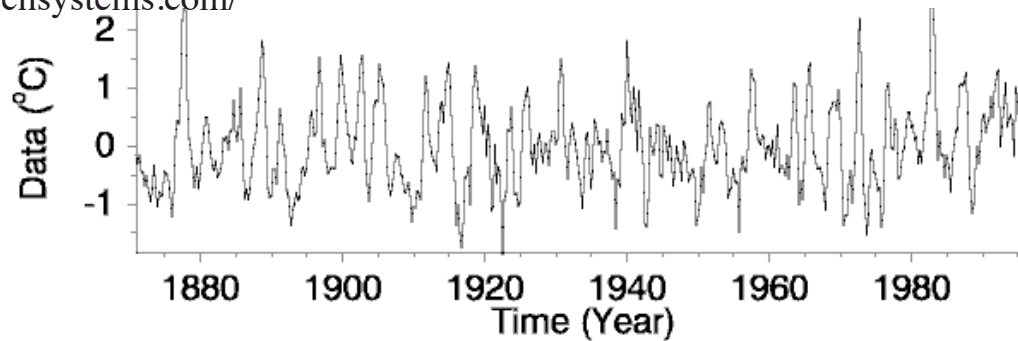
5 4 3 2 1 0

Wavelet packets



Analyse de la température de surface d'El Nino par ondelettes

<http://ion.researchsystems.com/>



(a) Data. (b) The wavelet power spectrum. The contour levels are chosen so that 75%, 50%, 25%, and 5% of the wavelet power is above each level, respectively. (c) The global wavelet power spectrum. Reference: Torrence, C. and G. P. Compo, 1998: A Practical Guide to Wavelet Analysis. Bull. Amer. Meteor. Soc., 79, 61-78.

Wigner-Ville transform

■ Search for a transform with following properties

Instantaneous density $T_x(t, f) \longrightarrow$ Energy $E = \int \int T_x(t, f) dt df$

Marginal densities

$$\int_{-\infty}^{\infty} T_x(t, f) dt = |X(f)|^2$$
$$\int_{-\infty}^{\infty} T_x(t, f) df = |x(t)|^2$$

■ Choice of a transform with these properties

Instantaneous energy density: $R(t, \tau) = x\left(t + \frac{\tau}{2}\right) x^*\left(t - \frac{\tau}{2}\right)$

Wigner-Ville transform: $W_x(t, f) = \int_{-\infty}^{\infty} R(t, \tau) e^{-j2\pi f\tau} d\tau = \mathcal{F}_{\tau}(R(t, \tau))$

Wigner-Ville transform

Bilinear transform $W_{x+y}(t, f) = W_x(t, f) + W_y(t, f) + 2 \operatorname{Re}(W_{xy}(t, f))$

Problem: interference terms without meaning

$$W_{xy}(t, f) = \int_{-\infty}^{\infty} x\left(t + \frac{\tau}{2}\right) y^*\left(t - \frac{\tau}{2}\right) e^{-j2\pi f\tau} d\tau$$

Improvement : filtering in the spectrum plane (or equivalent windowing in the time plane)

Pseudo Wigner-Ville $PW_x(t, f) = \int_{-\infty}^{\infty} p(\tau) x\left(t + \frac{\tau}{2}\right) x^*\left(t - \frac{\tau}{2}\right) e^{-j2\pi f\tau} d\tau$

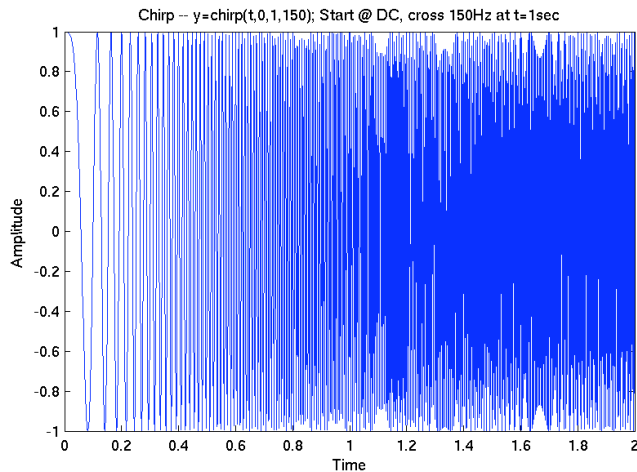


less interferences, but less resolution

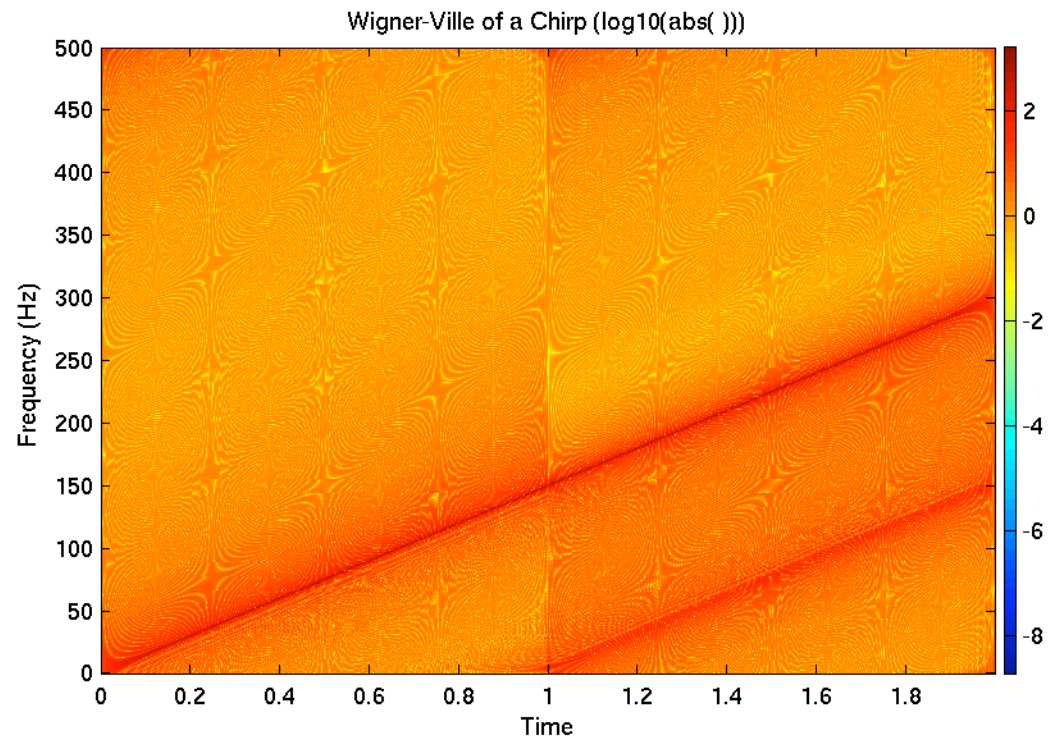
- Good properties for spectral analysis of frequency modulated signals (dolphins, whales, bat,...)

Wigner-Ville transform

■ Example : chirp



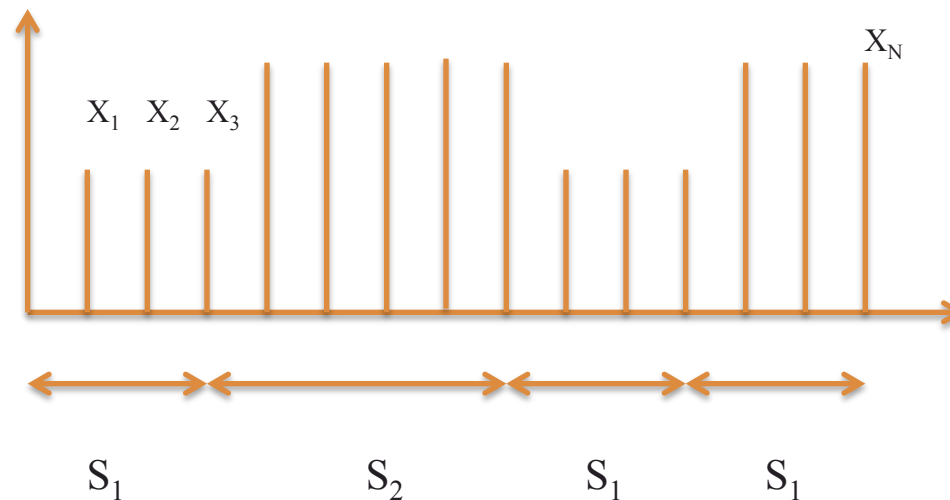
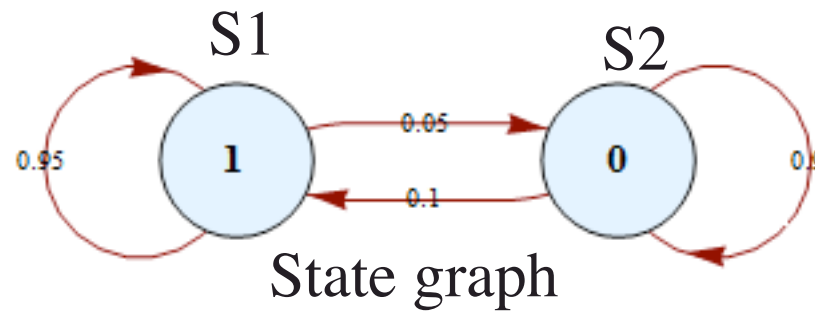
Linear frequency
modulation of a sine
wave



Wigner-Ville transform

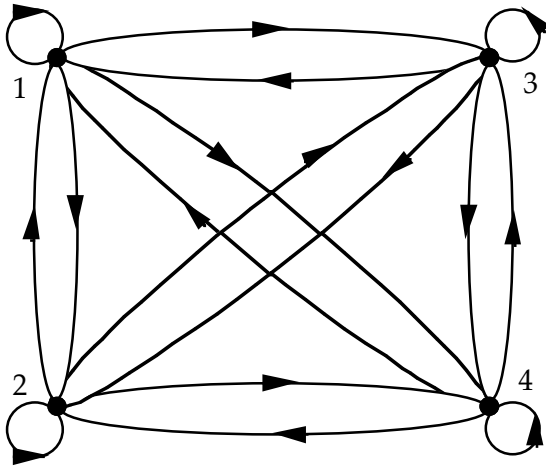
Hidden Markov models

- **Markov chain : sequence of states driven by transition probabilities (example of 2 states)**



Hidden Markov models

■ Markov chain : fully connected (example of 4 states)



$$A = \begin{bmatrix} a_{11} & a_{12} & a_{13} & a_{14} \\ a_{21} & a_{22} & a_{23} & a_{24} \\ a_{31} & a_{32} & a_{33} & a_{34} \\ a_{41} & a_{42} & a_{43} & a_{44} \end{bmatrix}$$

$$et, \forall (i, j), a_{ij} > 0$$

State transition matrix

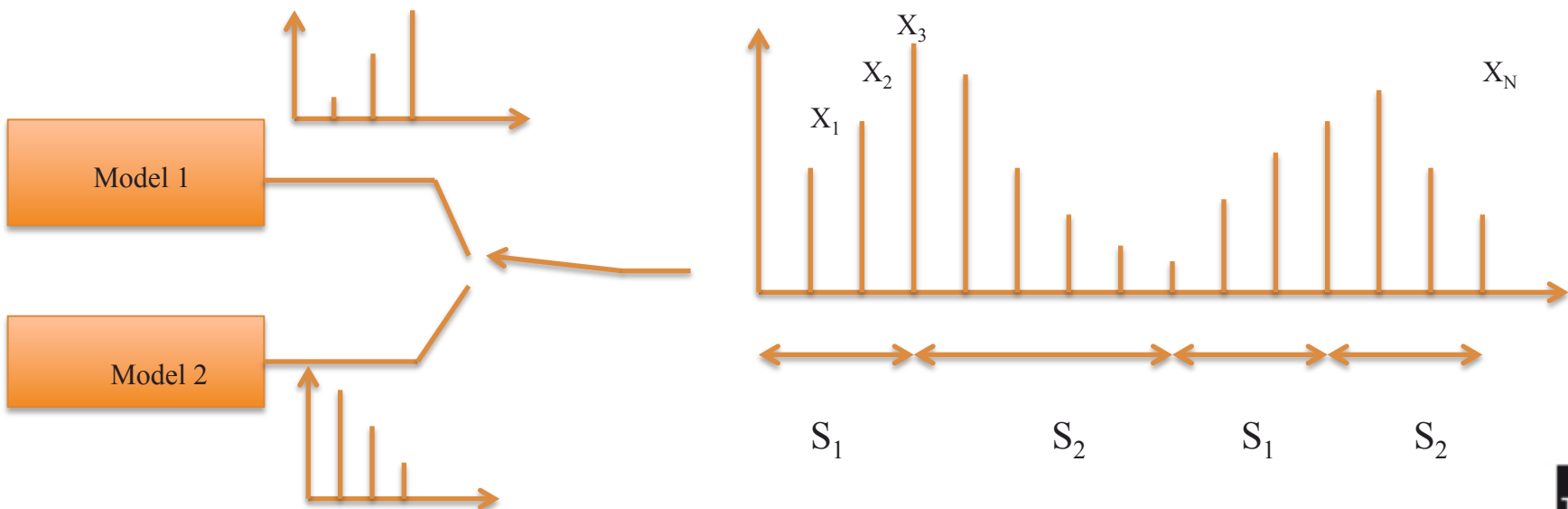
■ Markov chain description (M states)

- State set $\mathbf{s}=(S_i, i=1\dots M)$ where S_i means state i .
- Transition matrix $\mathbf{A} = (a_{ij} / i=1,\dots,M ; j=1,\dots,M)$ where a_{ij} is the probability of going from S_i at time $t-1$ to state S_j at time t .
- Initial state probabilities $\boldsymbol{\pi} = (\pi_i / i=1,\dots,M)$, where π_i is the probability of being in state S_i at time $t=0$.

Hidden Markov models



Hidden Markov chain

- No direct observation of states, but sequences of observations
- Description by $(\mathbf{s}, \mathbf{A}, \mathbf{p}, \pi)$
 - Markov chain description $(\mathbf{s}, \mathbf{A}, \pi)$
 - Probability distribution of observations in each state : $\mathbf{p} = (p_i(x) / i=1, \dots, M)$ where $p_i(x)$ is the probability density of observation x conditionally to state i .

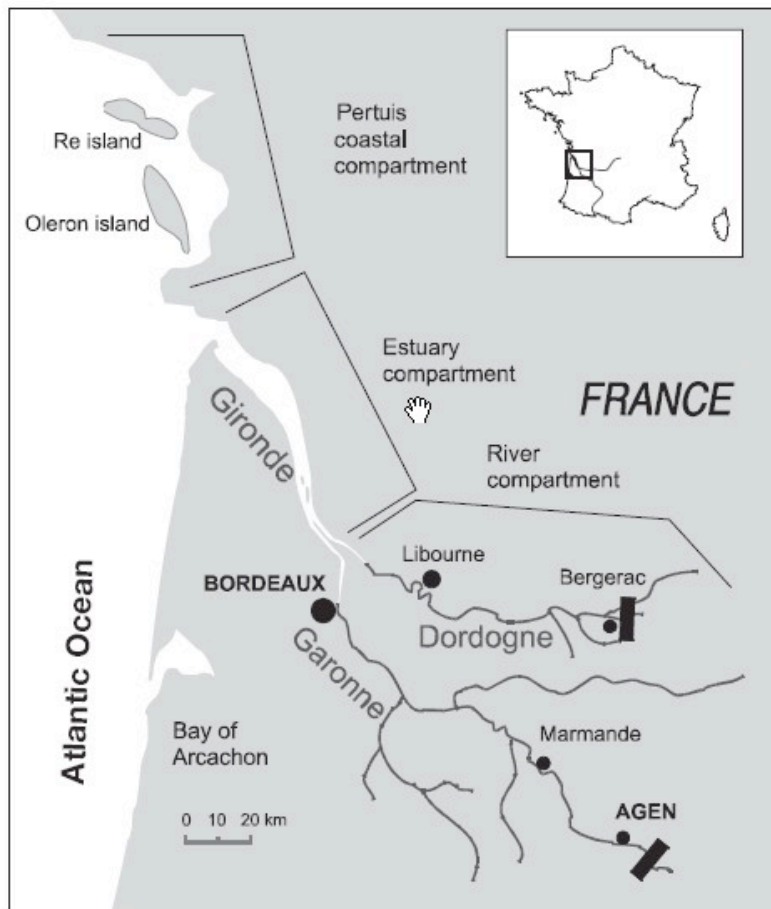


Hidden Markov models

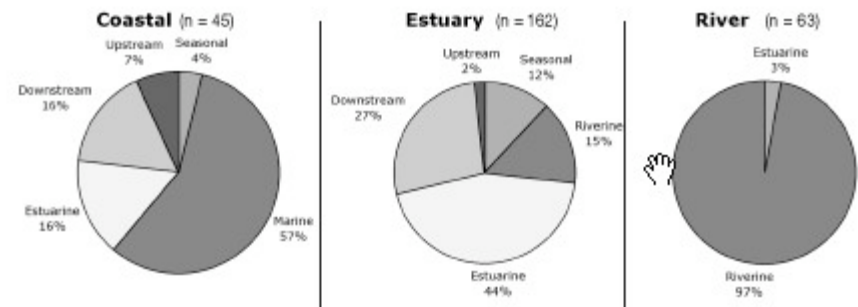
■ Problems to be solved

- **Pattern recognition problem:** Choice of the best model to describe time series (each model being associated to a pattern)
- Solution: maximization of the probability distribution $p(\mathbf{x}_{1:N}/\lambda_i)$ where $\mathbf{x}_{1:N}$ is the observation sequence x_1, x_2, \dots, x_N and $\lambda_i = (A_i, \mathbf{p}_i, \pi)$ is the model i with $i=1, \dots, M$
- **Classification problem:** recovering of the state sequence $\mathbf{s}_{1:N}$ from the noisy observation sequence $\mathbf{x}_{1:N}$
- Solution: maximization of the probability distribution $p(\mathbf{s}_{1:N}/\mathbf{x}_{1:N}, \lambda)$
-  Viterbi algorithm (or dynamic programming)
- **Estimation problem** : estimation of the model parameters
- Solution : maximization of the likelihood function $p(\mathbf{x}_{1:N}/\lambda)$
-  Baum-Welch algorithm

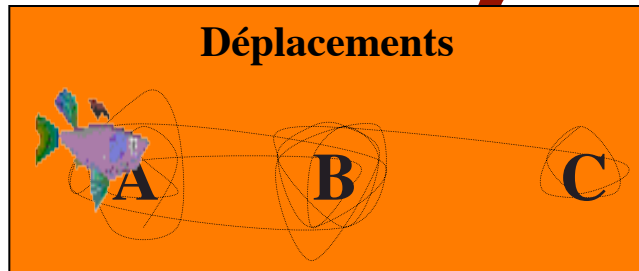
Migration of the eels



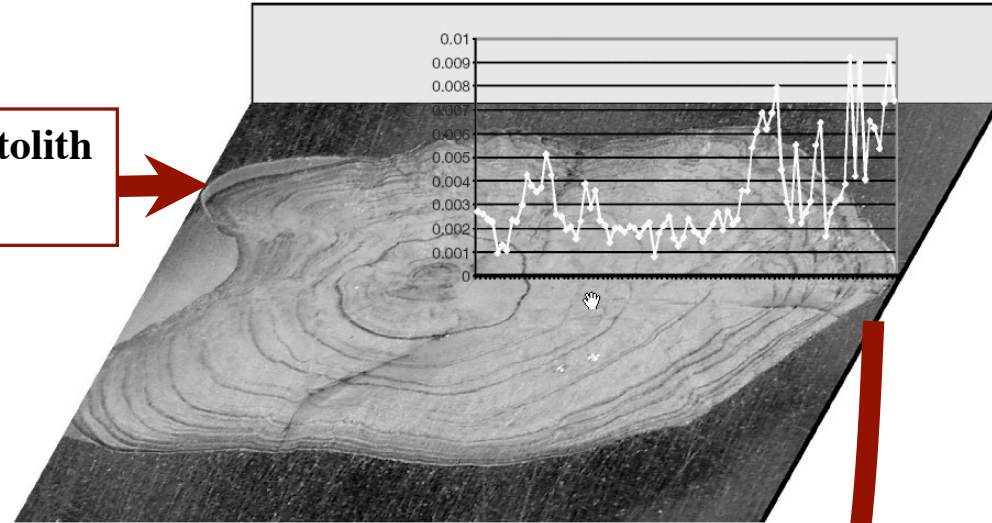
- Correlation between the ratio Strontium/Calcium of otoliths and migration areas (coastal, estuary, river)



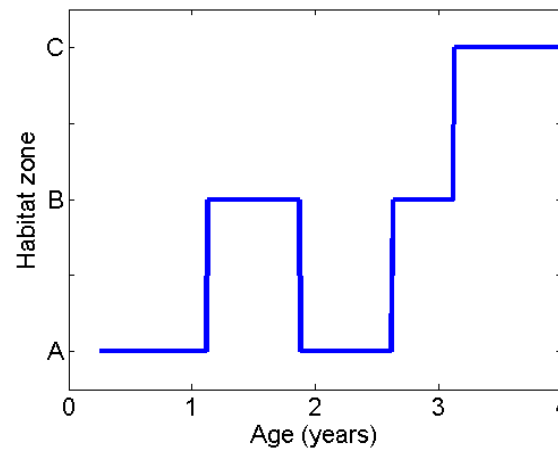
Migration of the eels



Signal from otolith analysis



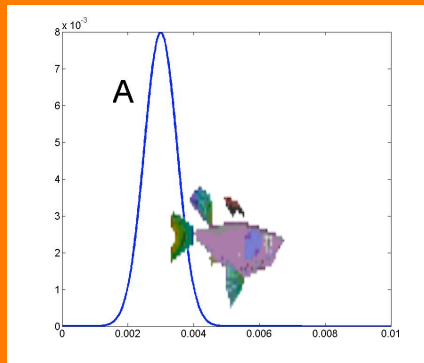
How to find the hidden states ?



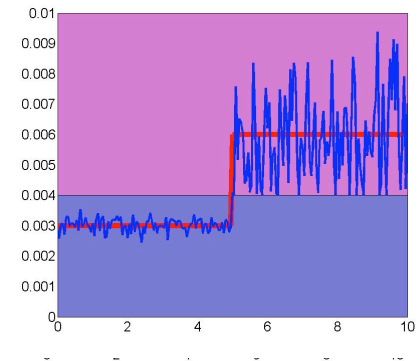
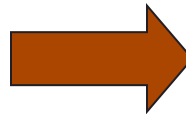
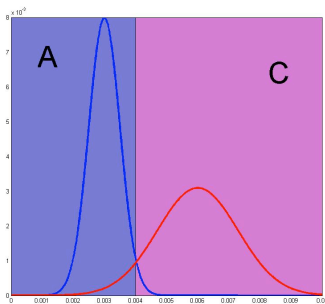
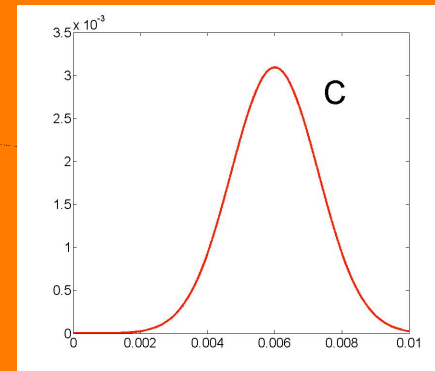
Hidden states : migration areas

Migration of the eels

Migrations



B

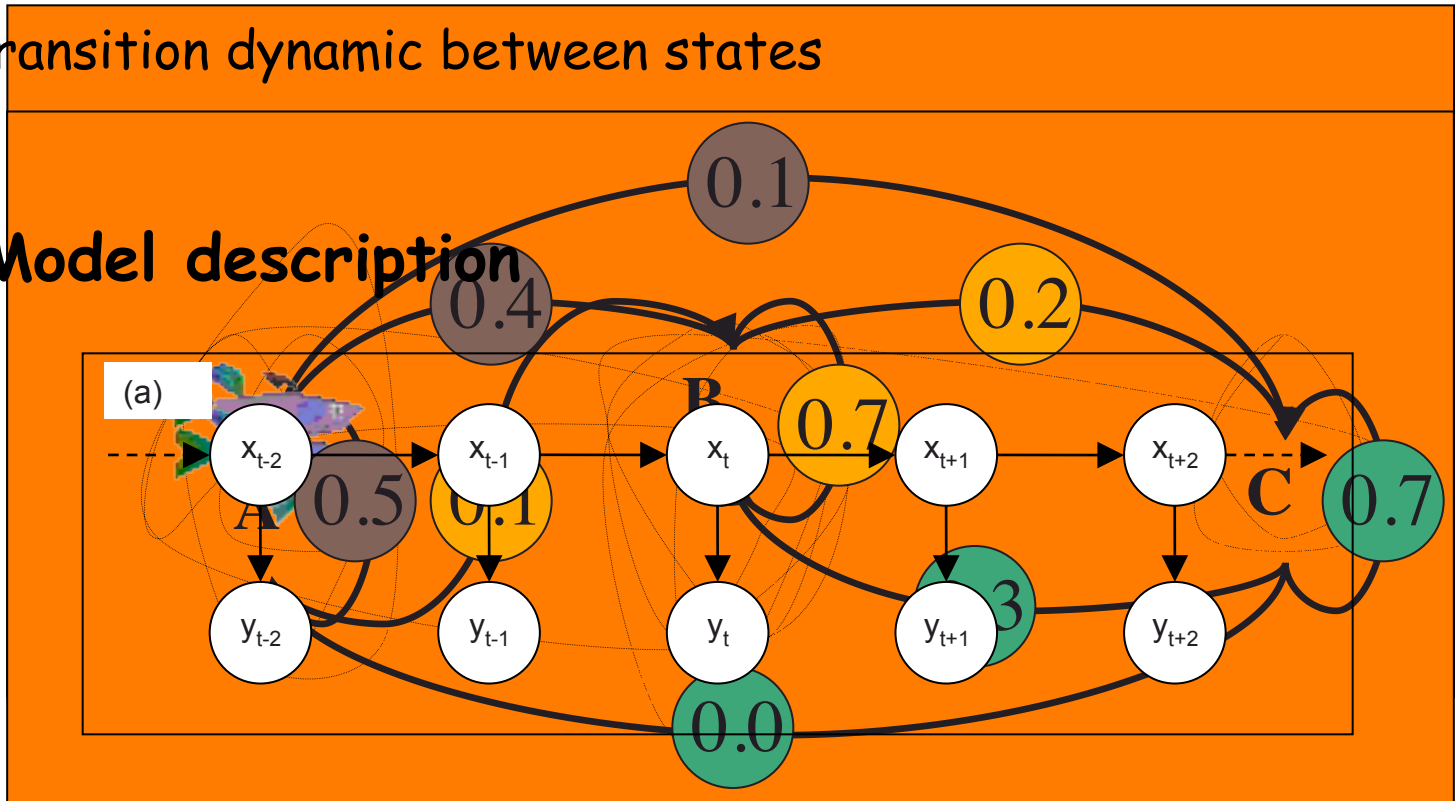


Migration of the eels

• HMM model characteristics

- 1 : probability distributions of the observations for each state
- 2 : transition dynamic between states

• Model description



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