A Gaussian state-space model for wind fields in the North-East Atlantic

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SWGEN - 2012

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Plan

Motivation

2 Context and goals

3 Wind data

4) Mode

- Gaussian Linear State-Space Model
- Parameter Estimation
- 5) Validation of the model
 - Parameters interpretation
 - Validation of the model
- Conclusion and extension of the model

Motivations for the use of weather generators

Many natural phenomena and human activities depend on wind conditions

- Production of electricity by wind turbines
- Evolution of a coast line
- Maritime transport
- Drift of objects in the ocean

• ...

Wind data generally available on short periods of time

- 50 years of data maximum
- Not enough to compute reliable estimates of the probability of complex events

Stochastic model used to simulate artificial wind conditions

Monte-Carlo methods

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Available data:

 Reanalysis data from ECMWF: wind speed intensity

Goals:

- to propose a stochastic state-space model for wind fields
- to generate realistic wind conditions

Context and goals

One of the main difficulties and goals is: reproducing space-time motions of meteorological systems *e.g.* propagation of low pressure systems.





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Wind data

Data under study: reanalysis data ERA Interim from ECMWF

Available with regular sampling: $\Delta x = 0.75^{\circ}$, $\Delta t = 6h$

To eliminate seasonality: study of months of January from 1979 to 2011

Wind speed are squared root transformed to reduce skewness

18 gridded points under study in the North-East Atlantic channel



Wind data

Some time series



Time shifts are observable on time series

Wind data

Some statistics computed on data



Figure: Spatial correlation at temporal lag 0 and autocorrelation function at location 5

location

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Model

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Modelling by a Gaussian state-space model [DK01], [CMR]

$$\left(egin{array}{rcl} X_t&=&bX_{t-1}+\epsilon_t\ Y_t(1)&=&Z(1)X_t+\eta_t(1)\ dots&dots&dots\ Y_t(K)&=&Z(K)X_t+\eta_t(K) \end{array}
ight. ext{for }t\geq 0,$$

 $Y_t \in \mathbb{R}^{K}$: vector of observations, K: number of locations studied $X_t \in \mathbb{R}$: scalar hidden state, ϵ and η are serially independent Gaussian white noises $\epsilon_t \sim \mathcal{N}(0, Q)$ and $\eta_t \sim \mathcal{N}(0, R)$

Parameters and equations interpretation:

- X : mean squared root wind conditions at regional scale, • temporal dynamic is contained in the state equation,
- Y: mean corrected squared root wind speed at local scale,
- Z : links regional and local scales,
- R : contains a part of the spatial covariance between studied locations.

$$\begin{cases} X_t = bX_{t-1} + \epsilon_t \\ Y_t(1) = Z(1)X_t + \eta_t(1) \\ \vdots & \vdots \\ Y_t(K) = Z(K)X_t + \eta_t(K) \end{cases} \text{ for } t \ge 0,$$

 $\epsilon_t \sim \mathcal{N}(\mathbf{0}, \mathcal{Q}) \; \mathsf{and} \; \eta_t \sim \mathcal{N}(\mathbf{0}, \mathcal{R})$

Stationarity:

Under the assumption |b| < 1, the autoregressive process X is stationary and so is the process Y [BD].

Identifiability:

Q is fixed in order to ensure identifiability of the model (study of the second order structure of the Gaussian process Y [TA10]).

Cross-correlation

The estimate of cross-correlation shows that some locations have temporal advance on others:



To capture this phenomenon, we add more flexibility to the model

More flexible dependance from regional mean wind speed

Model

To get a more flexible temporal dependance of observations from the regional mean wind speed, we consider the followed observation equation:

$$Y_t = Z_0 X_t + Z_1 X_{t-1} + \eta_t,$$

Taking $\tilde{X}_t = \begin{pmatrix} X_t \\ X_{t-1} \end{pmatrix}$, the system can be rewritten as: $\begin{cases} \tilde{X}_t &= \tilde{B}\tilde{X}_{t-1} + \tilde{\epsilon}_t \\ Y_t &= \tilde{Z}\tilde{X}_t + \eta_t \end{cases} \text{ for } t \ge 0,$ with $\tilde{B} = \begin{pmatrix} b & 0 \\ 1 & 0 \end{pmatrix}$, $\tilde{Z} = \begin{pmatrix} Z_0 & Z_1 \end{pmatrix}$ and $\tilde{\epsilon}_t \sim \mathcal{N}\left(\begin{pmatrix} 0 \\ 0 \end{pmatrix}, \begin{pmatrix} Q & 0 \\ 0 & 0 \end{pmatrix}\right)$.

This model is still a linear Gaussian state-space model.

Identifiability: This model brings new problems of identifiability....

Covariance structure: [Cre], [HR89], [Abr]

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(Observation equation: $Y_t = Z_0 X_t + Z_1 X_{t-1} + \eta_t$, $\eta_t \sim \mathcal{N}(0, R)$)

For parsimony purpose we choose to fit the covariance observation error matrix to a parametric family of spatial covariance:

Assume that R depends on the distance between locations:

$$R_{i,j} = \left(\sigma_i \sigma_j (\exp(-\lambda_1 d_{i,j}^2) + \lambda_2 \delta_{i,j})\right) \quad \text{for} \quad i,j \in \{1,...,K\},$$

th $(\sigma_1,...,\sigma_K,\lambda_1,\lambda_2)$ to be estimated and strictly positive.

Advantage: it reduces the number of parameters to estimate

Maximum Likelihood Estimation

Generalized Expectation-Maximization algorithm is used [CMR]:

- E-step: Linearity and Gaussianity of the model enable to use Kalman filter and smoother to derive the incomplete likelihood,
- M-step: Explicit forms of the parameters: \tilde{B} and \tilde{Z} ,
 - . *R* is estimated by numerical optimization of a part of the incomplete likelihood.

Starting point: Least square estimation on second order structure.

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Conclusion and extension of the mode

Estimate of *Z*:

Observation equation: $Y_t = Z_0 X_t + Z_1 X_{t-1} + \eta_t$,



Western locations depend more on X_t than on X_{t-1} and conversely for the eastern locations.

Estimate of *R*:

 $\text{Observation equation: } Y_t = Z_0 X_t + Z_1 X_{t-1} + \eta_t, \quad \eta_t \sim \mathcal{N}(0,R)$



Location 12 has the weakest error measurement, the most westerly and easterly locations have the largest ones.

Comparison of the smoothed hidden state \hat{X} to the squared root of wind speed Y at locations 1 and 18, with

$$\hat{X}_t = E(X_t | Y_1, ..., Y_T)$$



X is a good compromise between square root of wind speed at each location.

Marginal Distributions



Figure: Locations 1 and 12

Weak values are overestimated and high values are not correctly reproduced. Location 12 distribution matches more with the data distribution.

Temporal dynamic



• Time shifts are not entirely reproduced.

• For about one day time-lag, temporal dynamic is reproduced in some locations and underestimated for greater time-lags.

Spatial structure

10 - 0.9 - 0.8 - 0.7 - 0.6

Incation

Empirical correlation of the data





Spatial covariance of the observed process Y is underestimated but the shape is in part reproduced.

Spatial structure



Figure: Empirical correlation of Y at lag 0 against distance

Correlation is underestimated for distances smaller than 300km and variability is larger for distances greater than 400 km.

Comparison of different spatial structures



Figure: Corrrelation against distance for Gaussian variogram model and exponential variogram model

Spatial structure and others properties are not reproduced with the exponential model which overestimates some parameters and underestimates others. **Comparison of log-likelihood:**

Exponential variogram model 36779 Gaussian variogram model 43286

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Conclusion

Advantages of the model:

- Ease of interpretation of the parameters,
- Linearity and Gaussian properties: ease of implementation of estimation procedure
- · Some statistics are well reproduced by the model,

Drawbacks:

- Lack of flexibility of the model: it only catches average behaviors of wind speed in the area, does not account for the weather type,
- Linearity of the model may not be appropriate to catch possible local effects,

- Improving the temporal dynamic of this model,
- Adding Markovian regime switching to account for the weather type and to give more flexibility to the model,
- Modelling wind speed and direction at the same time will be more relevant: the dependance is different according the direction in which the wind is blowing, and we know that according the weather type the wind has prevailing direction.

P. Abrahamsen.

A review of gaussian random fields and correlation functions.

- Peter J. Brockwell and Richard A. Davis. Introduction to time series and forecasting. Springer Texts in Statistics. Second edition.
- Olivier Cappé, Eric Moulines, and Tobias Rydén. Inference in hidden Markov models. Springer Series in Statistics.

Noel A. C. Cressie.

Statistics for spatial data.

Wiley Series in Probability and Mathematical Statistics: Applied Probability and Statistics.

J. Durbin and S. J. Koopman.

Time series analysis by state space methods, volume 24 of *Oxford Statistical Science Series*.

Oxford University Press, Oxford, 2001.

J. Haslett and A.E. Raftery.

Space-time modelling with long-memory dependence: Assessing ireland's wind power ressource.

Applied Statistics, 1989.

Ailliot P. Tandeo, P. and E. Autret.

Linear gaussian state-space model with irregular sampling: application to sea surface temperature.

Stoch Environ Res Risk Assess, 2010.

diagonale *R* simulée: 0.57 0.56 0.56 0.55 0.56 0.56 0.55 0.55 0.57 0.59 0.62 0.57 0.63 0.63 0.65 0.59 0.59 0.63 diagonale *R* observée: 0.66 0.66 0.64 0.65 0.65 0.63 0.64 0.64 0.64 0.65 0.65 0.60 0.64 0.64 0.65 0.63 0.61 0.66